

# Geometry Understanding in Higher Dimensions

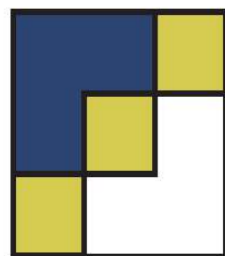
CHAIRE D'INFORMATIQUE ET SCIENCES NUMÉRIQUES

Collège de France - June 2017

## Statistics and Topological Data Analysis

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Laboratoire de  
Mathématiques  
Jean  
Leray

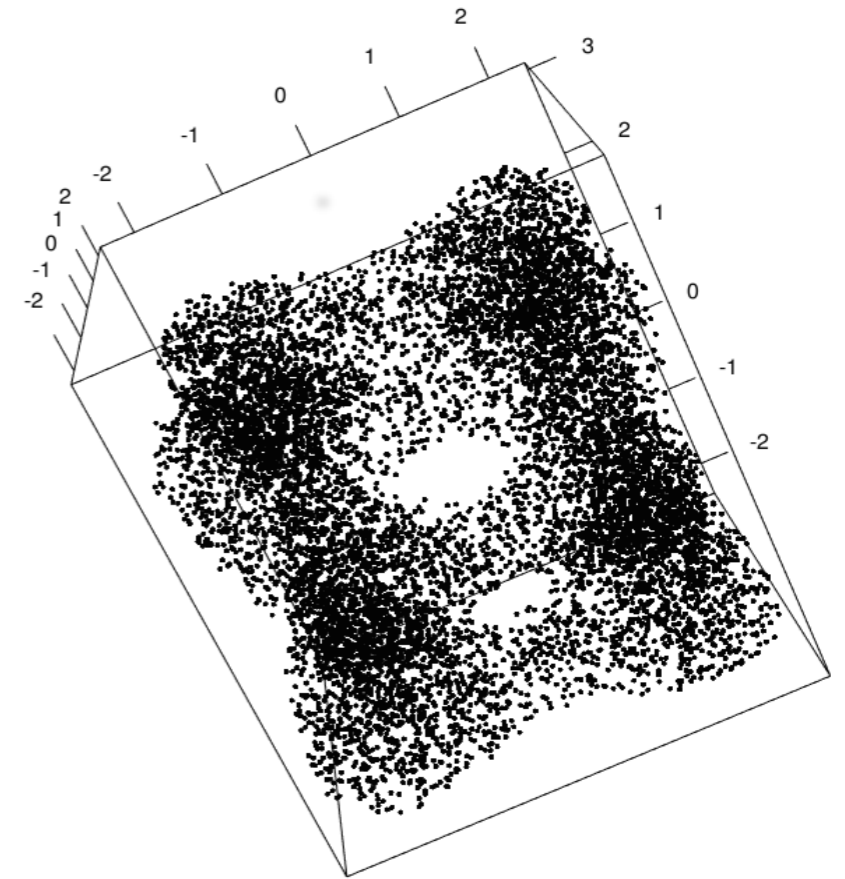
UMR 6629 - Nantes



# Introduction : Topological Data Analysis and Statistics

# Topological Data Analysis and Topological Inference

- The aim of TDA is to infer relevant qualitative and quantitative **topological structures** (clusters, holes ...) directly from the data.

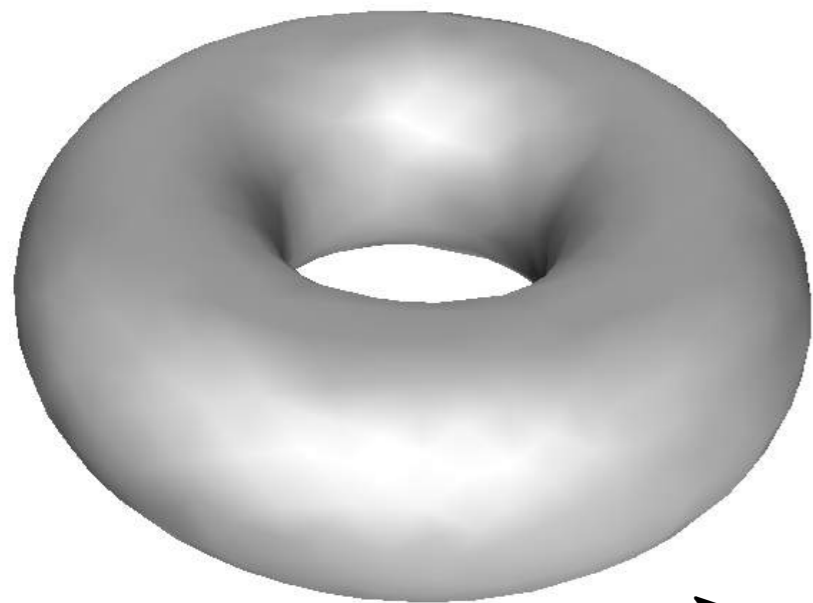


- data : typically point cloud  $X_n$
- Two popular methods in TDA : **Mapper algorithm** [Singh et al., 2007] and **persistent homology** [Edelsbrunner et al., 2002].

# Topological Data Analysis (TDA)

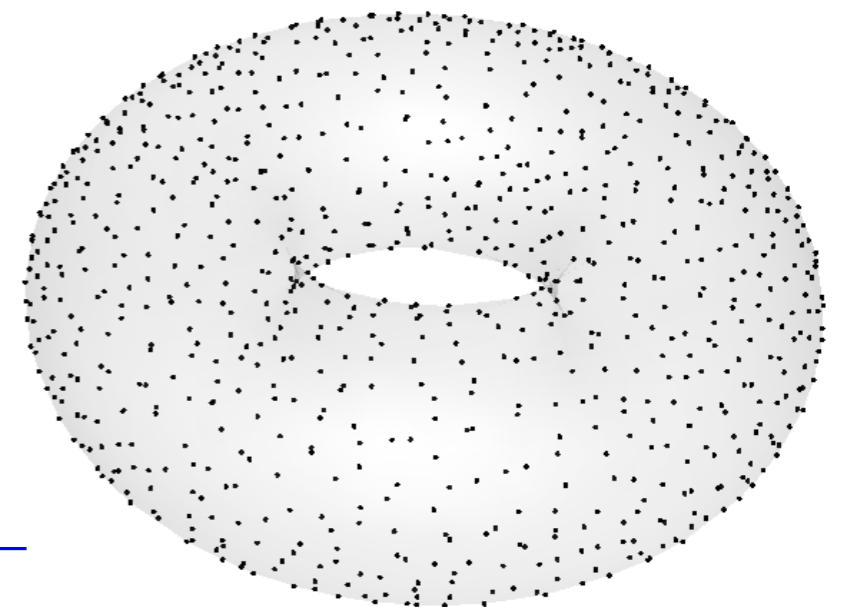
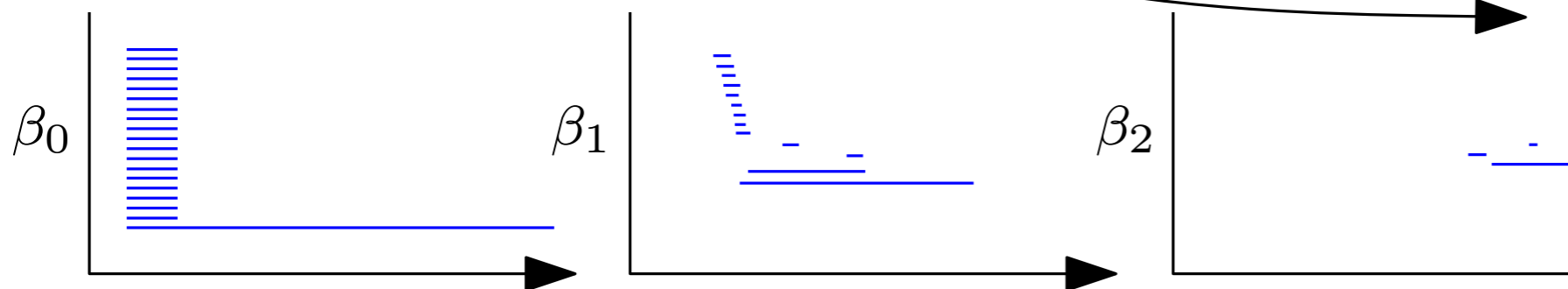
Why is topology interesting for data analysis?

- multiscale
- compact
- invariant under coordinate changes
- stable with respect to (small) perturbations
- informative



topological space

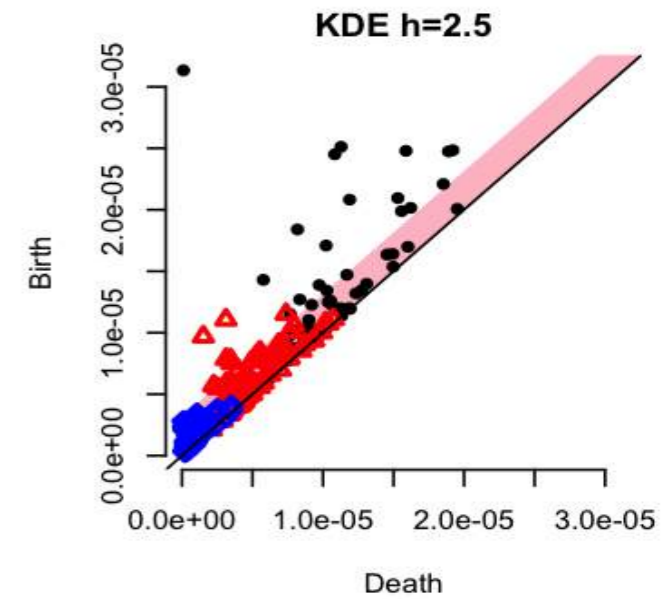
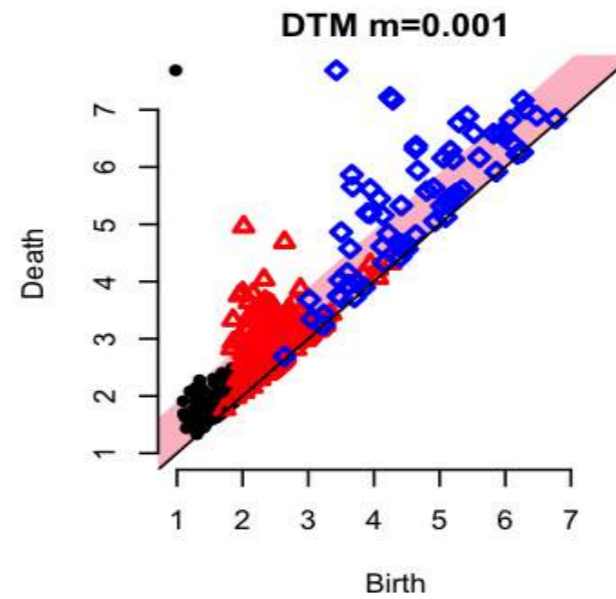
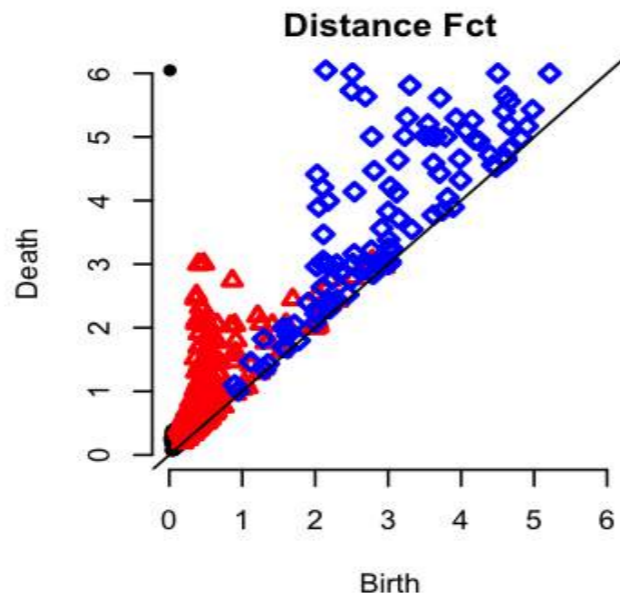
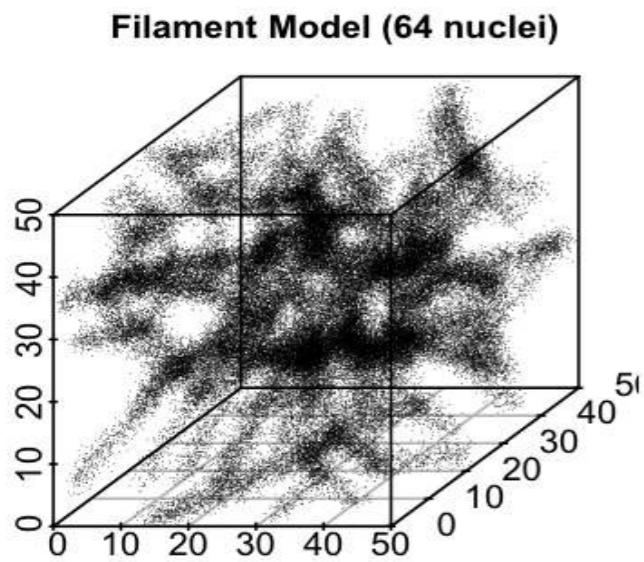
topological descriptors



point cloud

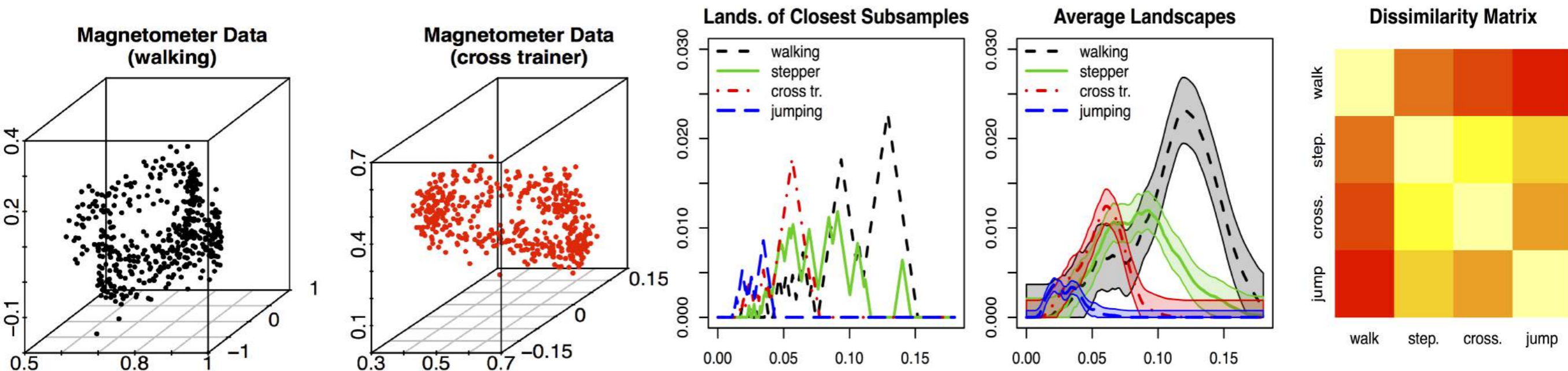
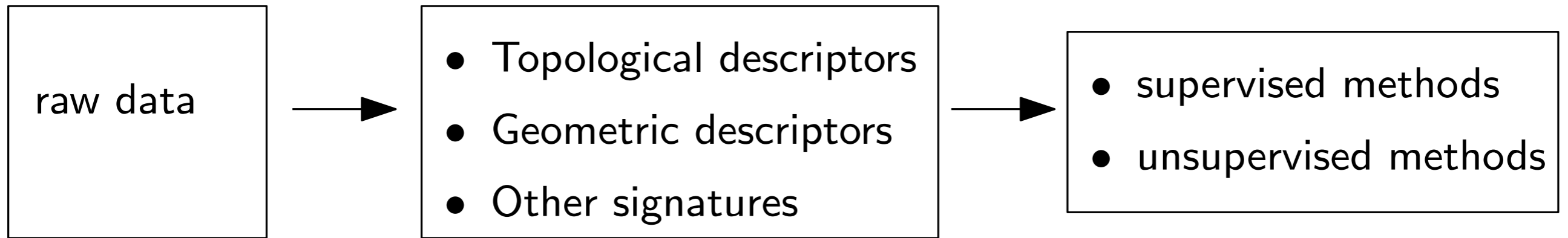
# Topological Data Analysis (TDA)

- For **exploratory analysis**, visualization



# Topological Data Analysis (TDA)

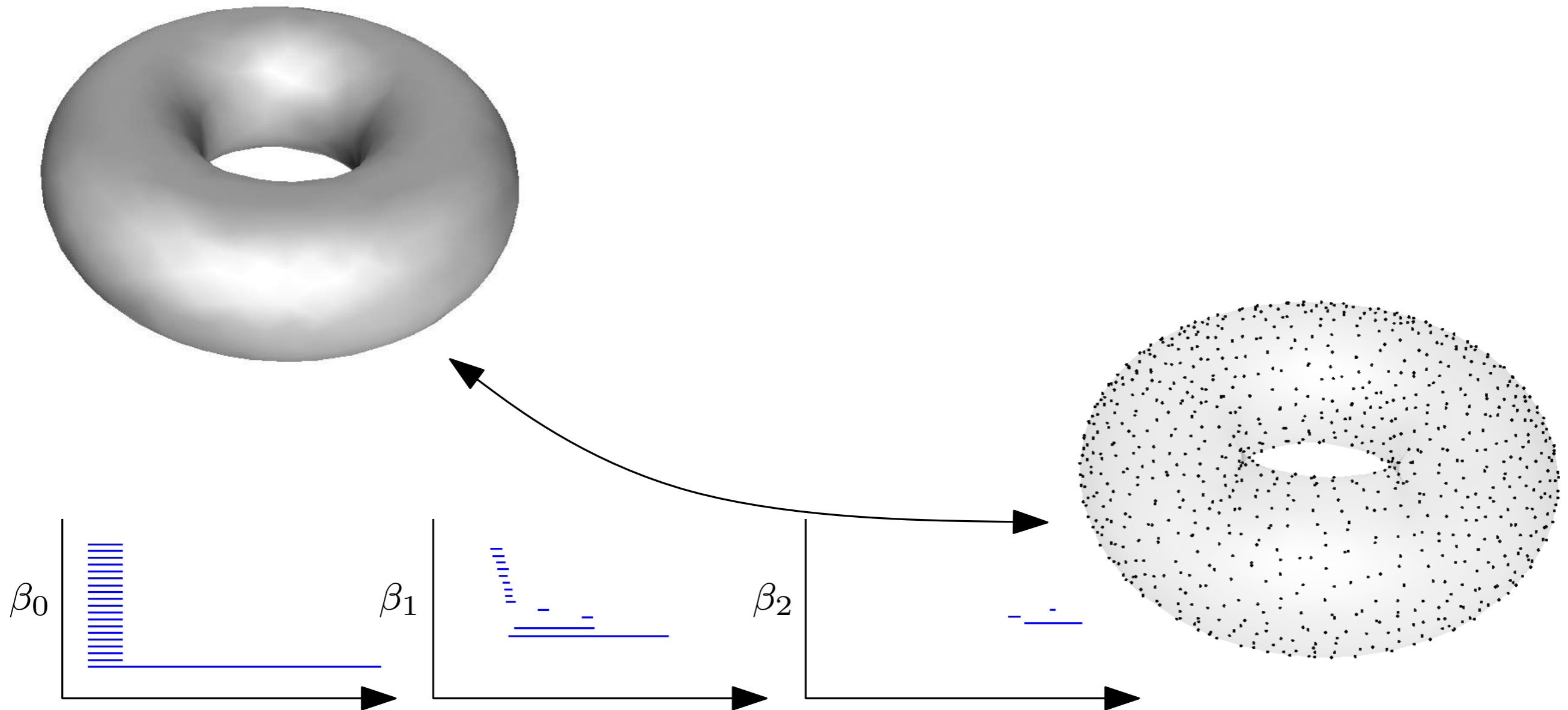
- For **exploratory analysis**, visualization
- For **feature extraction** and statistical learning



# Statistics, Learning and TDA

A **statistical approach to TDA** means that :

- we consider data as generated from an unknown distribution
- the inferred topological features by TDA methods are seen as estimators of topological quantities describing an underlying object.



# Statistics, Learning and TDA

Directions of research (non-exhaustive list):

- Consistency / convergence of TDA methods: [Chazal15 JMLR], [Bobrowski 17 Bernouilli]
- Confidence regions for TDA [Fasy 14 AoS] [Chazal 15 JOCG ]
- Central tendency for persistent homology [Turner 14 DCG] [Fasy15 Nips]
- Robust methods for TDA [Chazal 17, EJS Chazal 17 JMLR]
- Representations of persistence in Euclidean spaces [Bubenik15 JMLR] [Adams15]
- Develop kernels for topological descriptors [Reininghaus 15 IEEE] [Carriere 17 ICML ]
- Statistical analysis of Mapper [Carriere 17]
- ...



# Homology and Persistent homology

# Topological Stability and Regularity

Topological inference : under “regularity assumptions”, topological properties of  $X$  can be recovered from (the off-sets) of a close enough object  $Y$ .

# Topological Stability and Regularity

Topological inference : under “regularity assumptions”, topological properties of  $\mathbb{X}$  can be recovered from (the off-sets) of a close enough object  $\mathbb{Y}$ .

- The *local feature size* is a local notion of regularity :

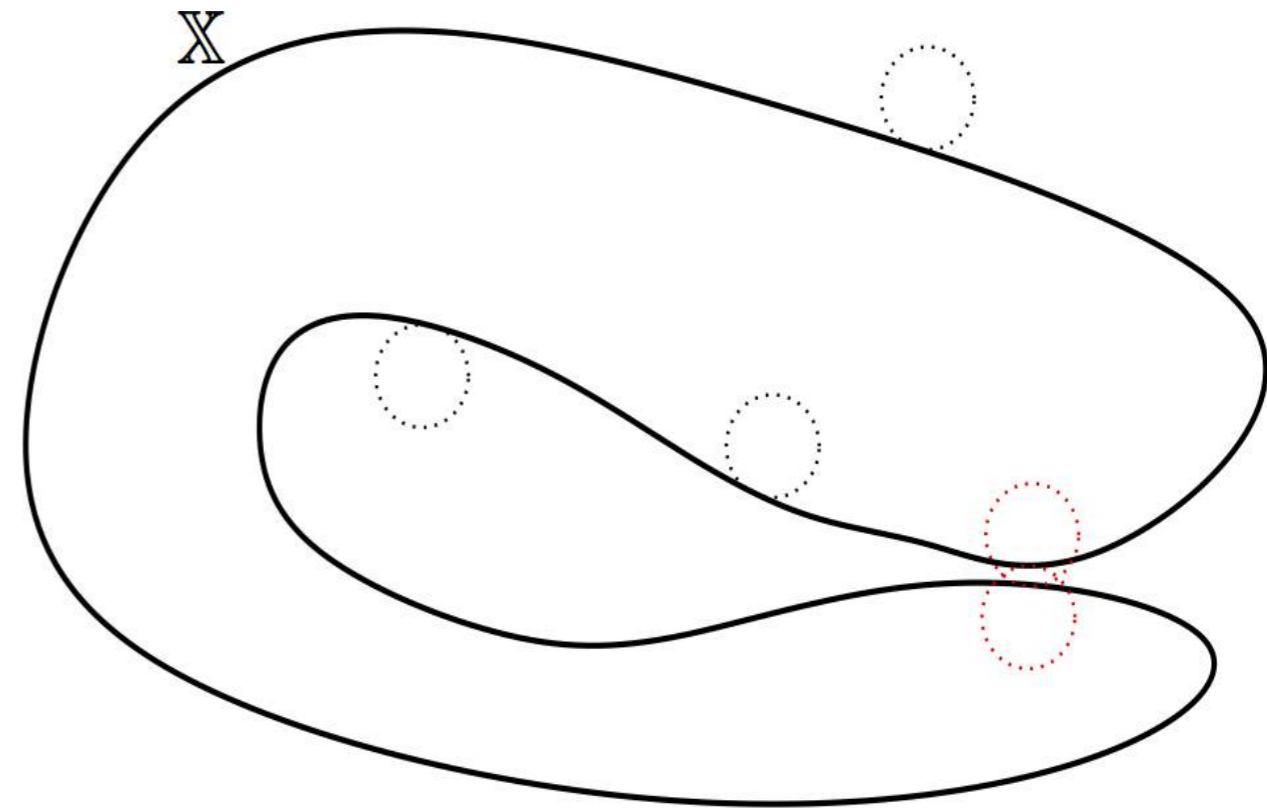
For  $x \in \mathbb{X}$ ,  $\text{lfs}_{\mathbb{X}}(x) := d(x, \mathcal{M}(\mathbb{X}^c))$ .

- The global version of the local feature size is the *reach* [Federer, 1959] :

$$\kappa(\mathbb{X}) = \inf_{x \in \mathbb{X}^c} \text{lfs}_{\mathbb{X}}(x).$$

The reach is small if either  $\mathbb{X}$  is not smooth or if  $\mathbb{X}$  is close to being self-intersecting.

- Weak feature size and its extensions [Chazal and Lieutier, 2007] (by considering the critical values of  $d_{\mathbb{X}}$ ).



# Topological Stability and Regularity

Topological inference : under “regularity assumptions”, topological properties of  $\mathbb{X}$  can be recovered from (the off-sets) of a close enough object  $\mathbb{Y}$ .

$$d_H(\mathbb{X}, \mathbb{Y}) = \inf \{ \alpha \geq 0 \mid \mathbb{X} \subset \mathbb{Y}^\alpha \text{ and } \mathbb{Y} \subset \mathbb{X}^\alpha \}$$

Example :

**Theorem** [Chazal and Lieutier, 2007]: Let  $\mathbb{X}$  and  $\mathbb{Y}$  be two compact sets in  $\mathbb{R}^d$  and let  $\varepsilon > 0$  be such that  $d_H(\mathbb{X}, \mathbb{Y}) < \varepsilon$ ,  $\text{wfs}(\mathbb{X}) > 2\varepsilon$  and  $\text{wfs}(\mathbb{Y}) > 2\varepsilon$ . Then for any  $0 < \alpha < 2\varepsilon$ ,  $\mathbb{X}^\alpha$  and  $\mathbb{Y}^\alpha$  are homotopy equivalent.

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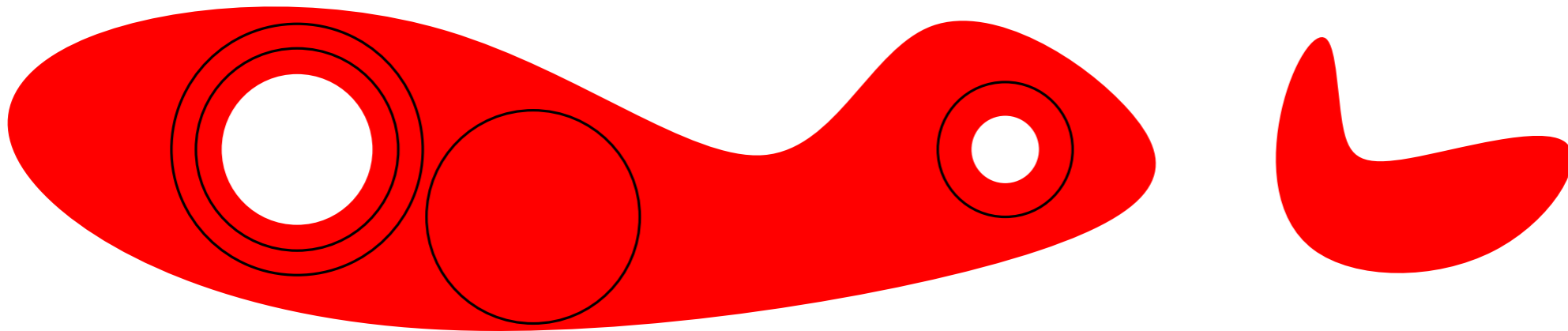
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Sampling conditions in Hausdorff metric.

Statistical analysis of homotopy inference can be deduced from support estimation of a distribution under the Hausdorff metric.

# Homology inference

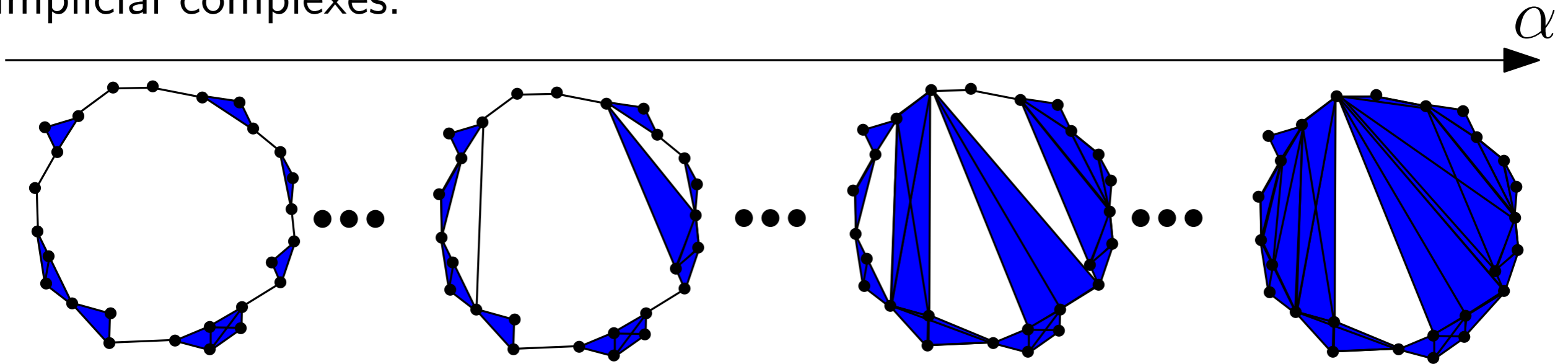
- **Homotopy** is not easy to compute in practice.
- **Singular homology** provides an algebraic description of “holes” in a geometric shape (connected components, loops, etc ...)
- **Betti number**  $\beta_k$  is the rank of the  $k$ -th homology group.
- **Computational Topology** : Betti numbers can be computed on simplicial complexes.



**Homology inference** [Niyogi et al., 2008 and 2011] [Balakrishnan et al., 2012] : The Betti number (actually the homotopy type) of Riemannian manifolds with positive reach can be recovered with high probability from offsets of a sample on (or close to) the manifold.

# Persistent homology

Starting from a point cloud  $\mathbb{X}_n$ , let  $\text{Filt} = (\mathcal{C}_\alpha)_{\alpha \in \mathcal{A}}$  be a filtration of nested simplicial complexes.

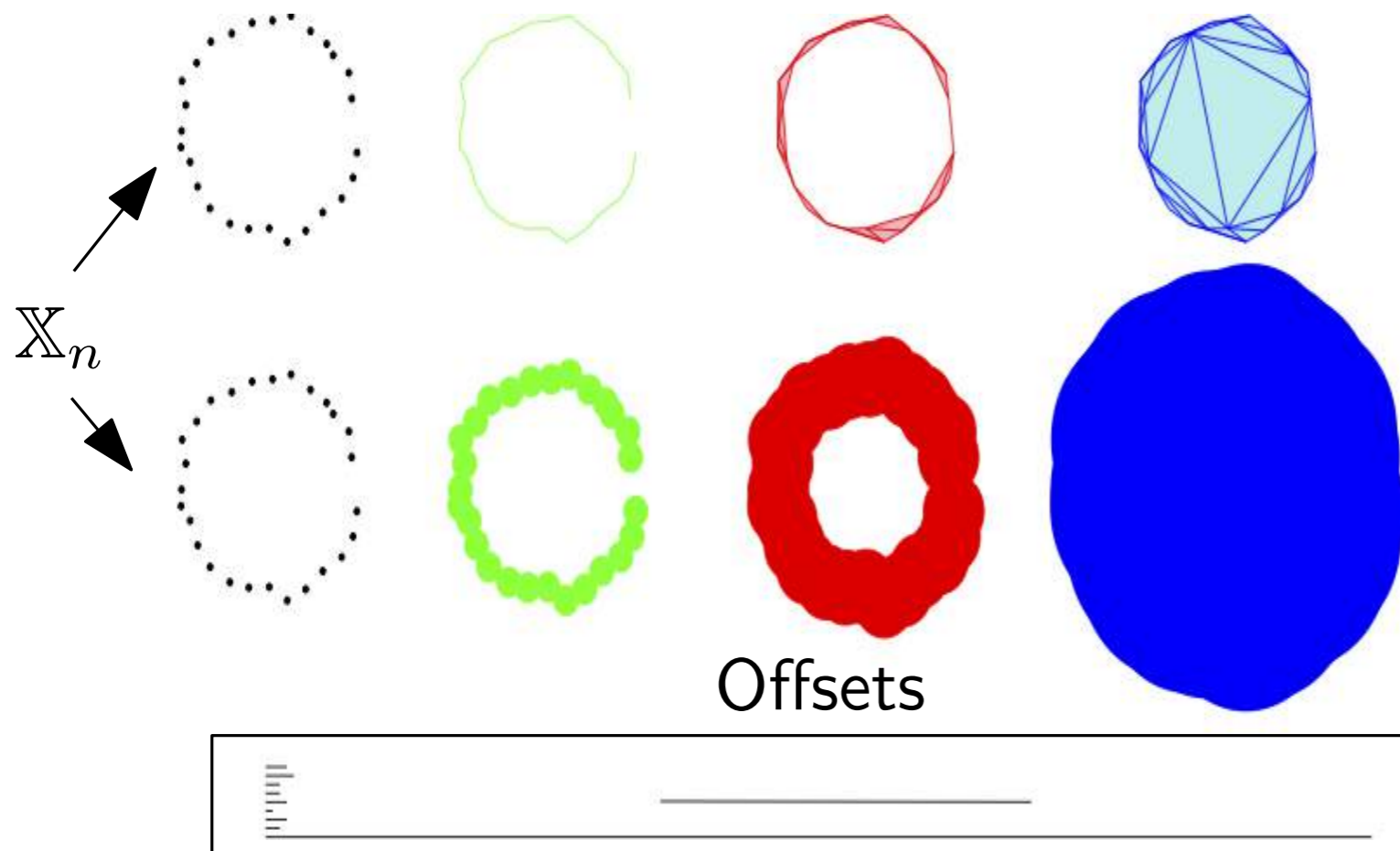


Persistent homology: identification of “persistent” topological features along the filtration.

- multiscale information ;
- more stable and more robust ;

# Barecodes and Persistence Diagrams

Filtration of simplicial  
complexes  $\text{Filt}(\mathbb{X}_n)$



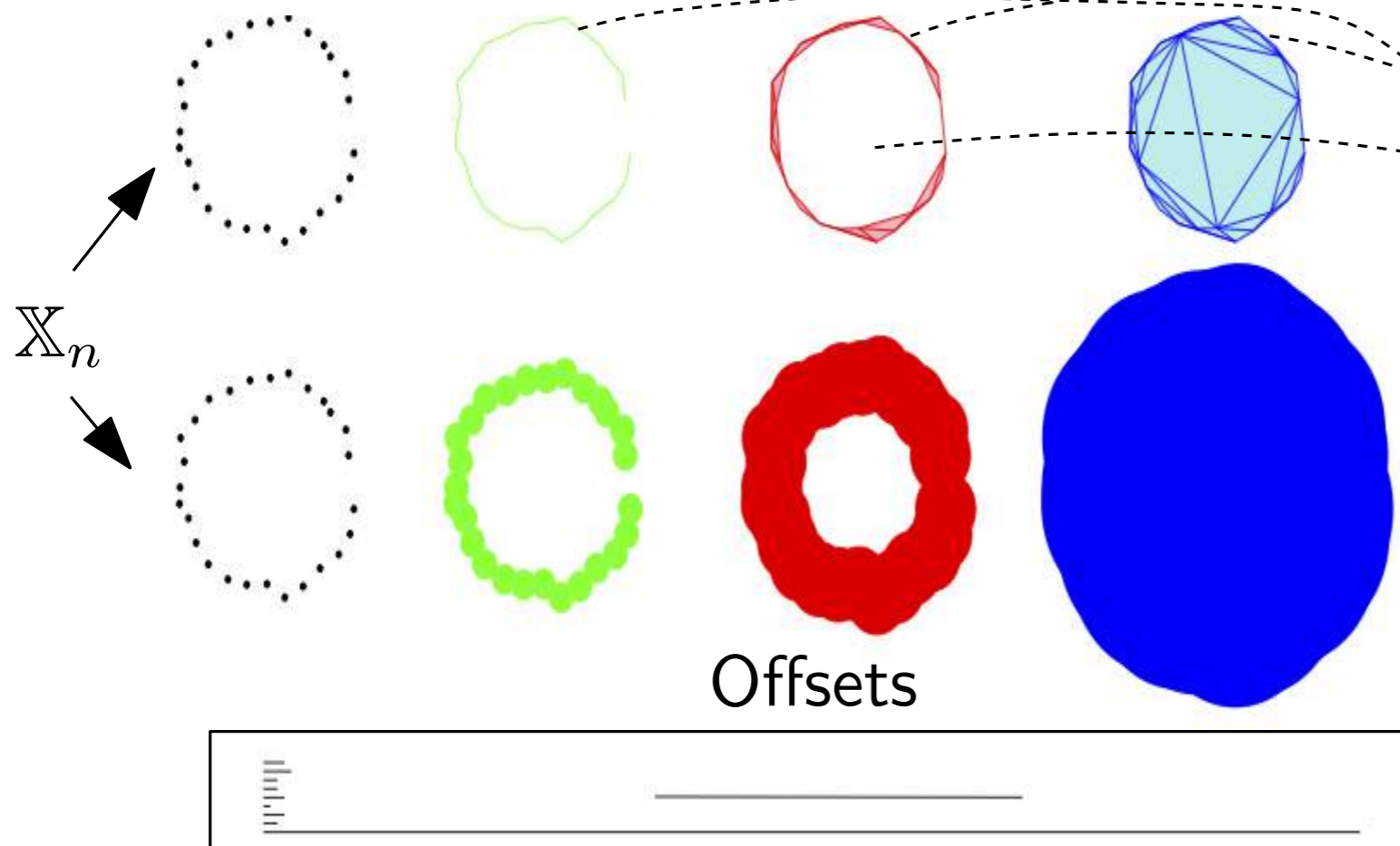
Offsets

Barcode



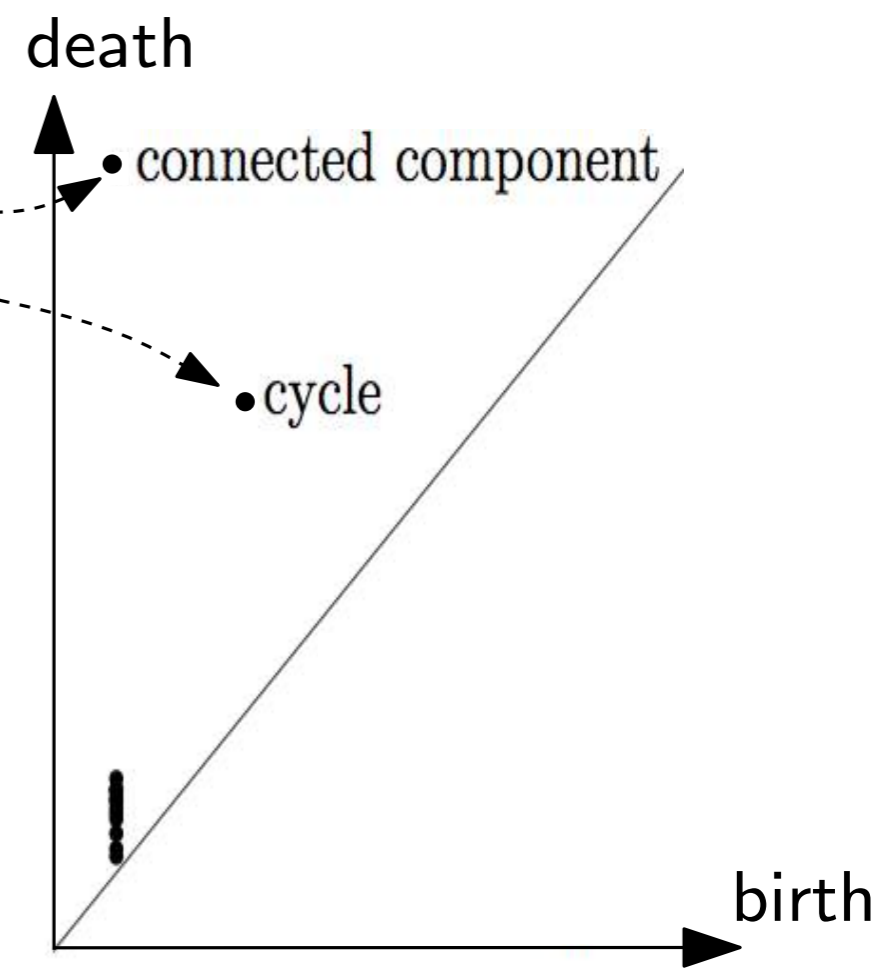
# Barecodes and Persistence Diagrams

Filtration of simplicial complexes  $\text{Filt}(\mathbb{X}_n)$



Offsets

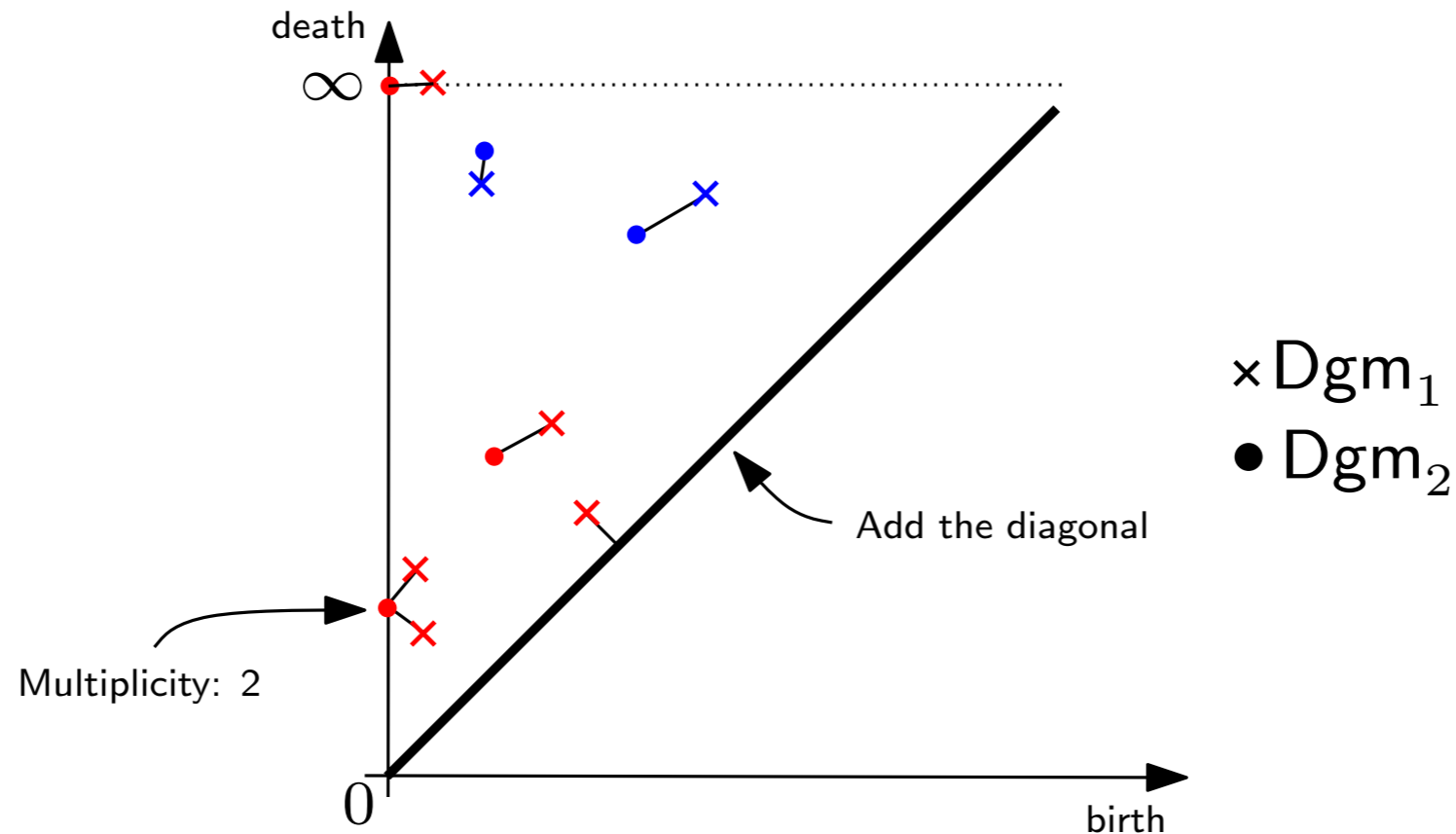
Barcode



$\text{Dgm}(\text{Filt}(\mathbb{X}_n))$

Persistence diagram of the filtration  $\text{Filt}(\mathbb{X}_n)$  built on  $\mathbb{X}_n$ .

# Distance between persistence diagrams and stability



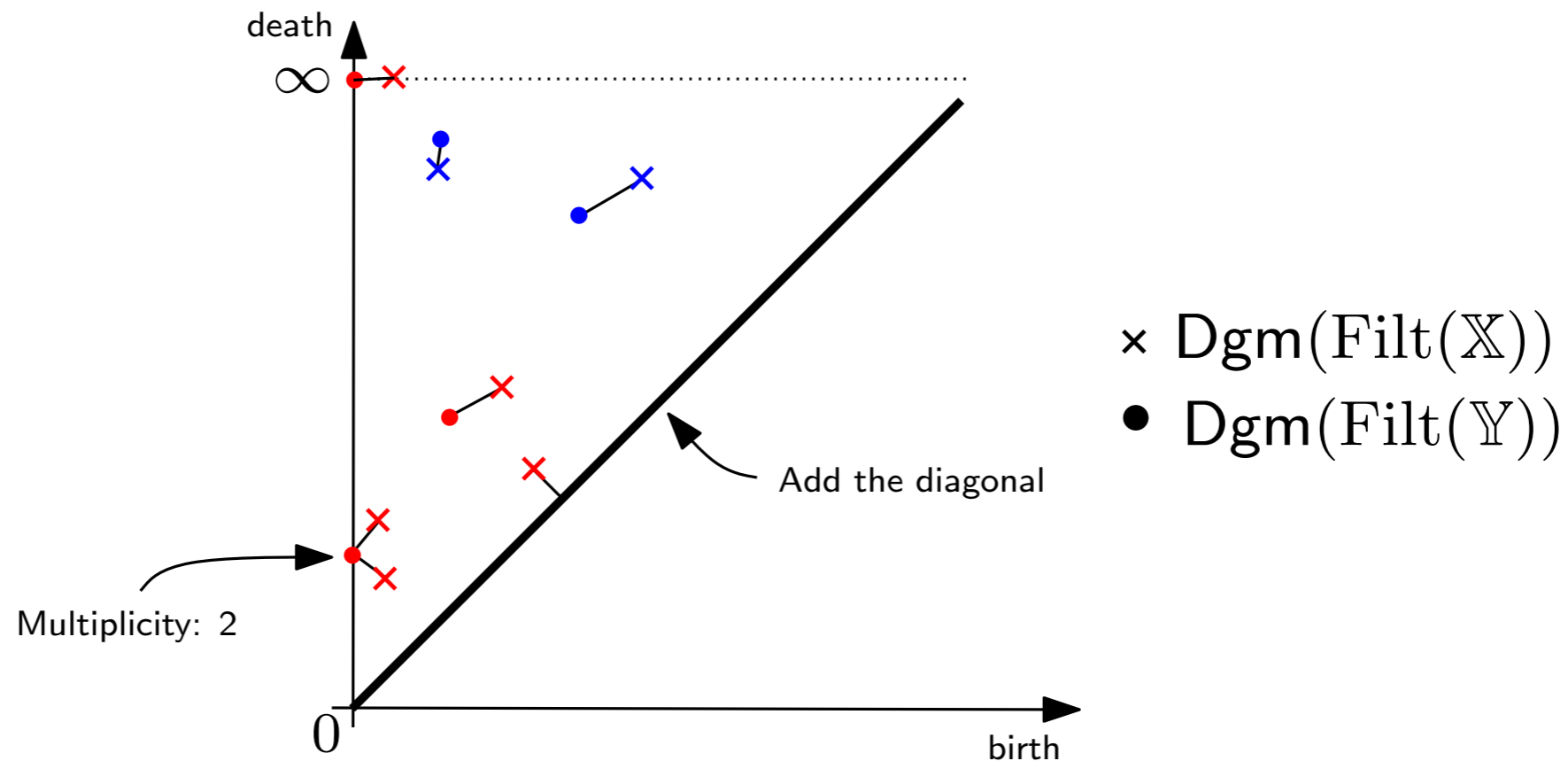
The **bottleneck distance** between two diagrams  $Dgm_1$  and  $Dgm_2$  is

$$d_b(Dgm_1, Dgm_2) = \inf_{\gamma \in \Gamma} \sup_{p \in Dgm_1} \|p - \gamma(p)\|_\infty$$

where  $\Gamma$  is the set of all the bijections between  $Dgm_1$  and  $Dgm_2$  and

$$\|p - q\|_\infty = \max(|x_p - x_q|, |y_p - y_q|).$$

# Distance between persistence diagrams and stability



**Theorem** [Chazal et al., 2012]: For any compact metric spaces  $(\mathbb{X}, \rho)$  and  $(\mathbb{Y}, \rho')$ ,

$$d_b (\text{Dgm}(\text{Filt}(\mathbb{X})), \text{Dgm}(\text{Filt}(\mathbb{Y}))) \leq 2 d_{\text{GH}} (\mathbb{X}, \mathbb{Y}).$$

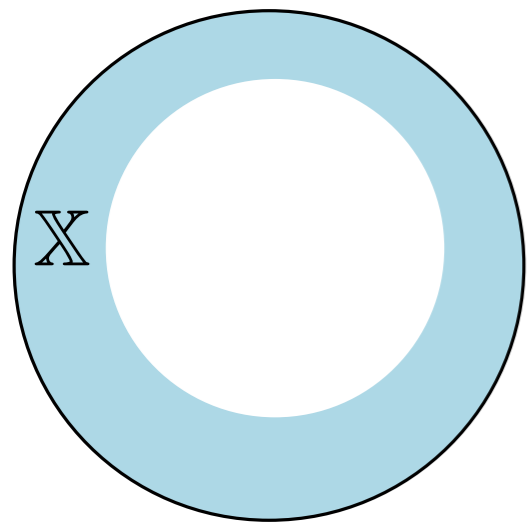
Consequently, if  $\mathbb{X}$  and  $\mathbb{Y}$  are embedded in the same metric space  $(\mathbb{M}, \rho)$  then

$$d_b (\text{Dgm}(\text{Filt}(\mathbb{X})), \text{Dgm}(\text{Filt}(\mathbb{Y}))) \leq 2 d_{\text{H}} (\mathbb{X}, \mathbb{Y}).$$

Statistics  
and  
Persistent homology

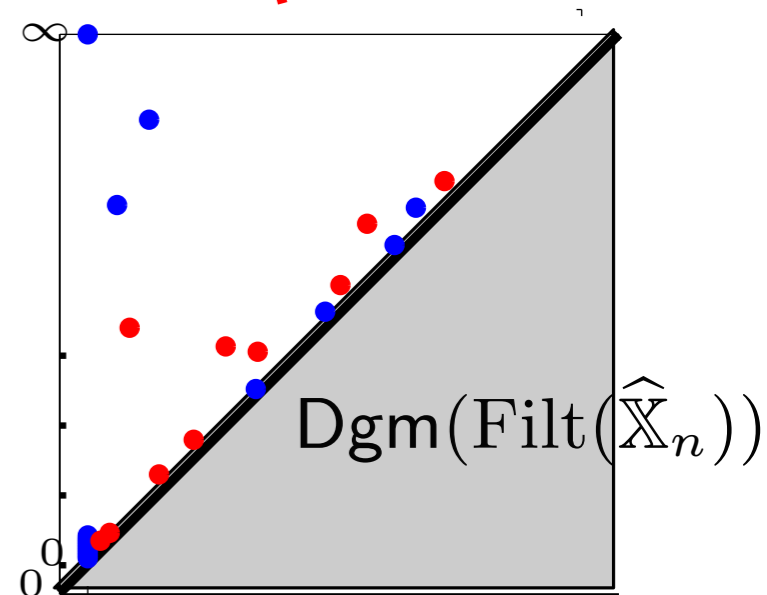
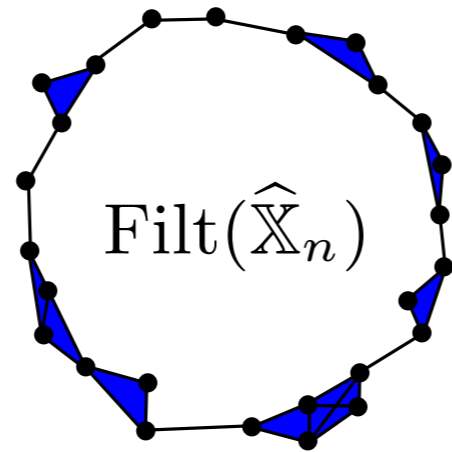
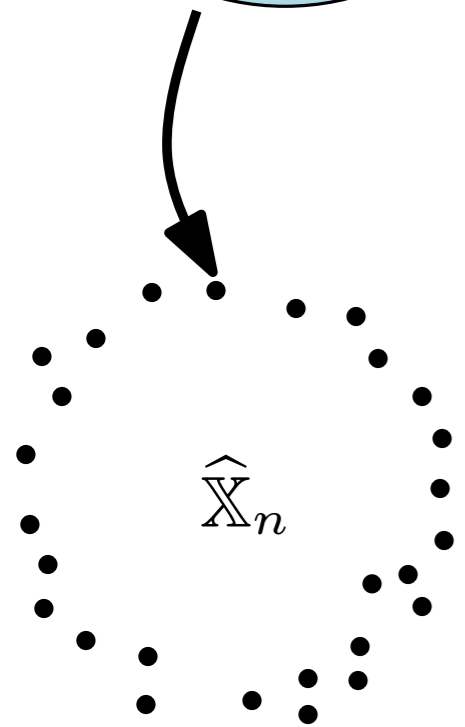
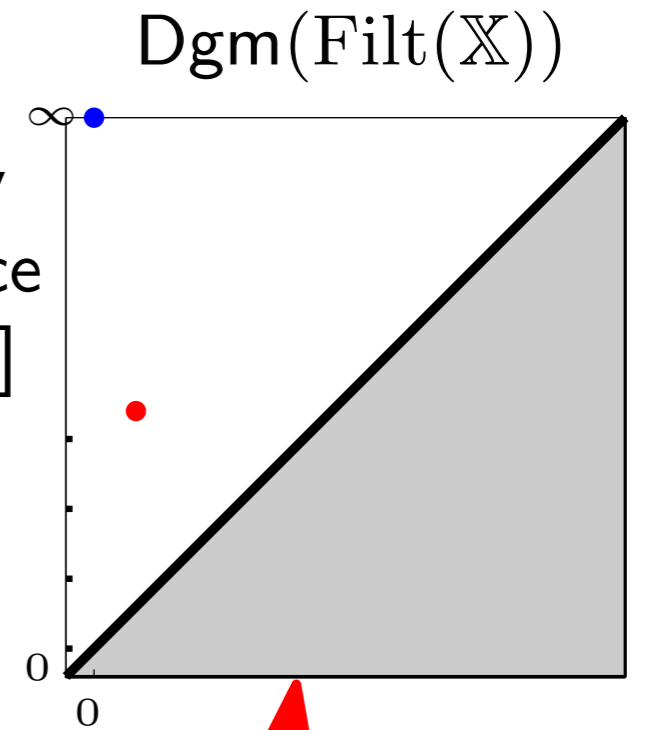
# Persistence diagram inference [Chazal 2015 JMLR]

$(M, \rho)$  metric space  
 $X$  compact set in  $M$ .



$\text{Filt}(X)$

well defined for any  
 compact metric space  
 [Chazal et al., 2012]



Convergence  
 ???

Estimator of  $Dgm(\text{Filt}(K))$

$n$  points sampled in  $X$   
 according to  $\mu$

# Persistence diagram inference

For  $a, b > 0$ ,  $\mu$  satisfies the  $(a, b)$ -standard assumption on its support  $\mathbb{X}_\mu$  if for any  $x \in X_\mu$  and any  $r > 0$  :

$$\mu(B(x, r)) \geq \min(ar^b, 1).$$

$\mathcal{P}(a, b, \mathbb{M})$  : set of all the probability measures satisfying the  $(a, b)$ -standard assumption on the metric space  $(\mathbb{M}, \rho)$ .

**Theorem:** For  $a, b > 0$  :

$$\sup_{\mu \in \mathcal{P}(a, b, \mathbb{M})} \mathbb{E} \left[ d_b(\text{Dgm}(\text{Filt}(\mathbb{X}_\mu)), \text{Dgm}(\text{Filt}(\widehat{\mathbb{X}}_n))) \right] \leq C \left( \frac{\ln n}{n} \right)^{1/b}$$

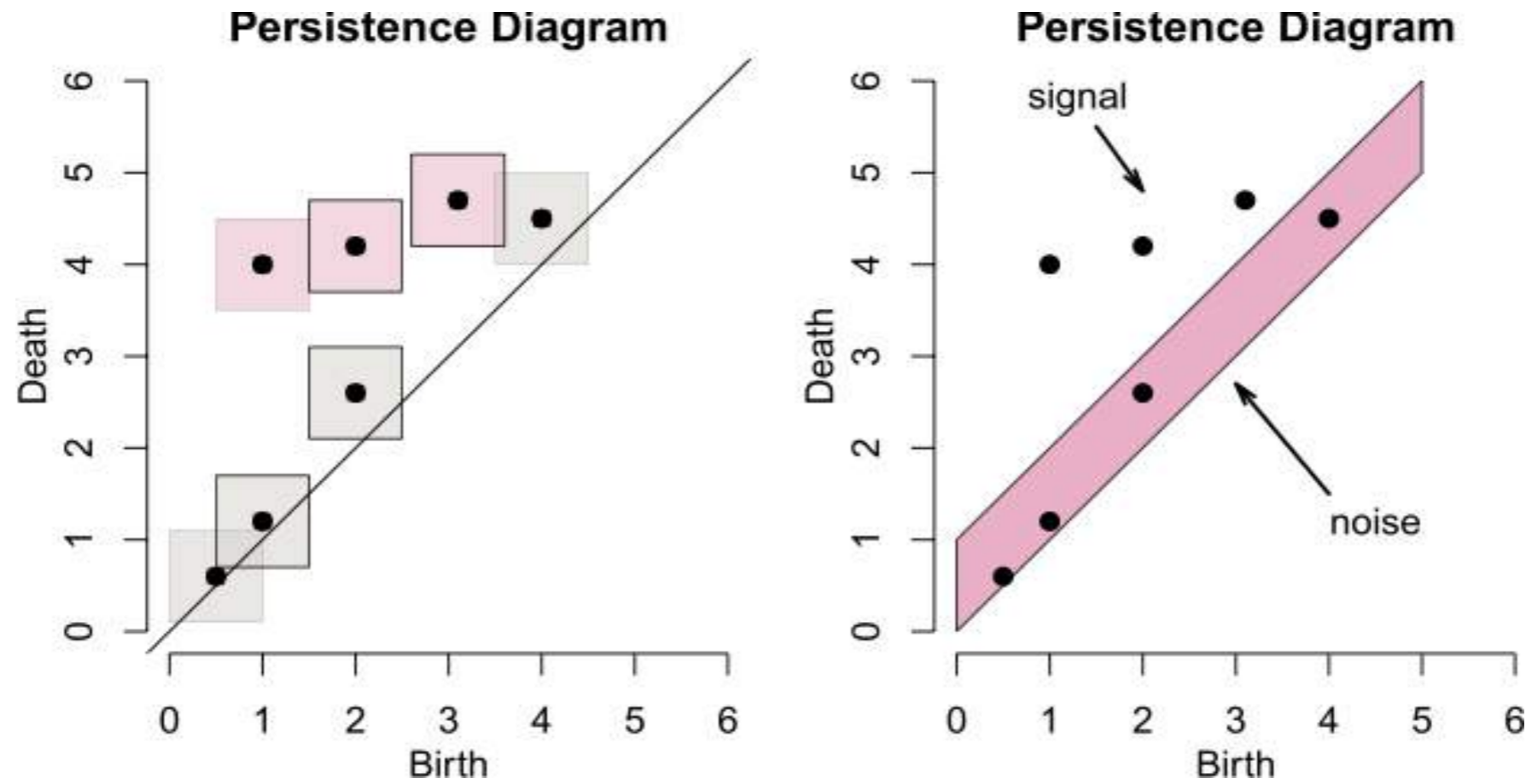
where  $C$  only depends on  $a$  and  $b$ .

Under additional technical hypotheses, for any estimator  $\widehat{\text{Dgm}}_n$  of  $\text{Dgm}(\text{Filt}(\mathbb{X}_\mu))$ :

$$\liminf_{n \rightarrow \infty} \sup_{\mu \in \mathcal{P}(a, b, \mathbb{M})} \mathbb{E} \left[ d_b(\text{Dgm}(\text{Filt}(\mathbb{X}_\mu)), \widehat{\text{Dgm}}_n) \right] \geq C' n^{-1/b}$$

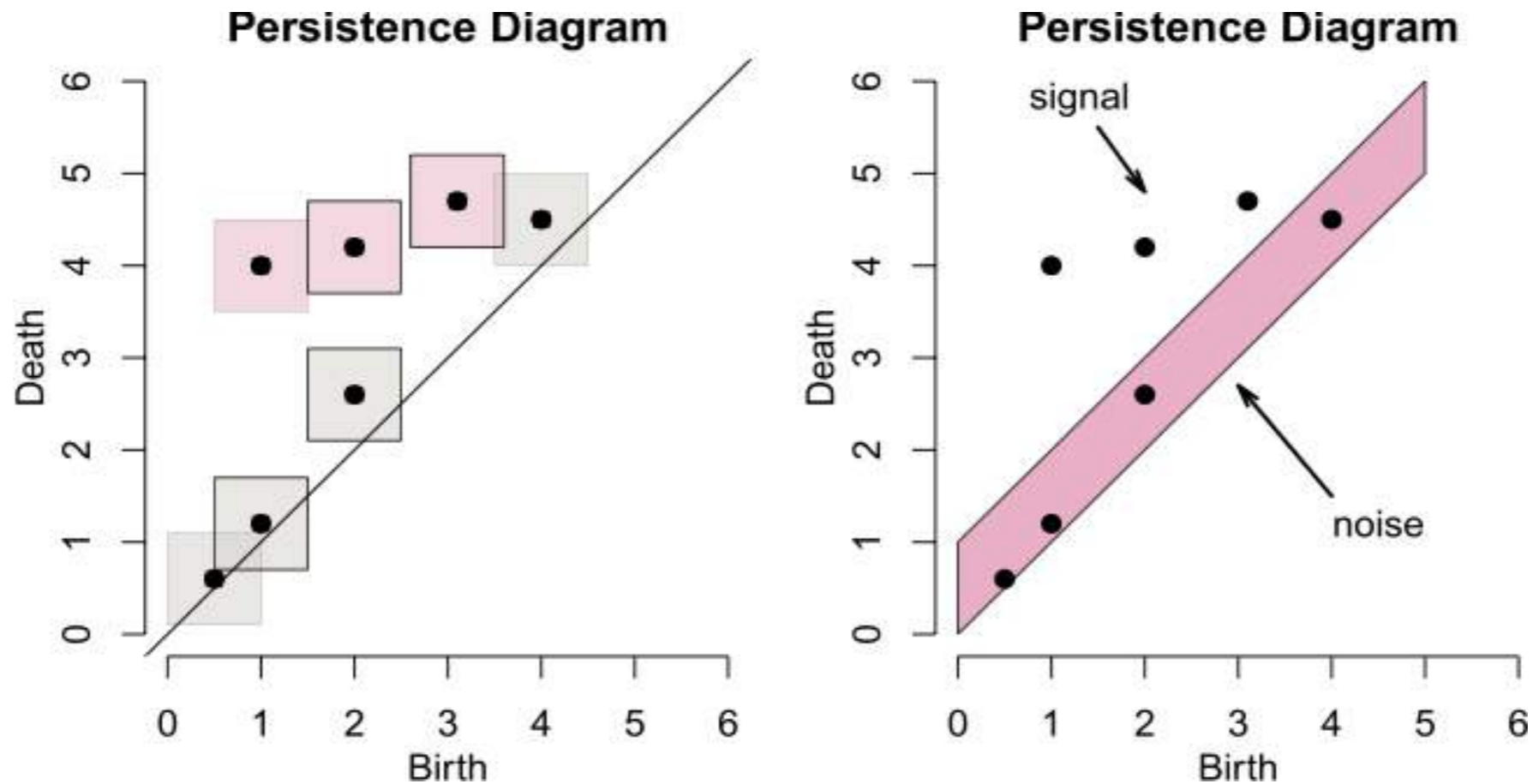
where  $C'$  is an absolute constant.

# Confidence sets for persistence diagrams [Fasy 2014 AoS]



$$P \left( \text{Dgm}(\text{Filt}(K)) \in \hat{\mathcal{R}} \right) \geq 1 - \alpha \quad ??$$

# Confidence sets for persistence diagrams [Fasy 2014 AoS]



$$P \left( \text{Dgm}(\text{Filt}(K)) \in \hat{\mathcal{R}} \right) \geq 1 - \alpha \quad ??$$

Using the Hausdorff stability, we can define confidence sets for persistence diagrams:

$$d_b \left( \text{Dgm}(\text{Filt}(K)), \text{Dgm}(\text{Filt}(\mathbb{X}_n)) \right) \leq d_H(K, \mathbb{X}_n).$$

It is sufficient to find  $c_n$  such that

$$\limsup_{n \rightarrow \infty} \left( d_H(K, \mathbb{X}_n) > c_n \right) \leq \alpha.$$



# Confidence sets for persistence diagrams [Fasy 2014 AoS]

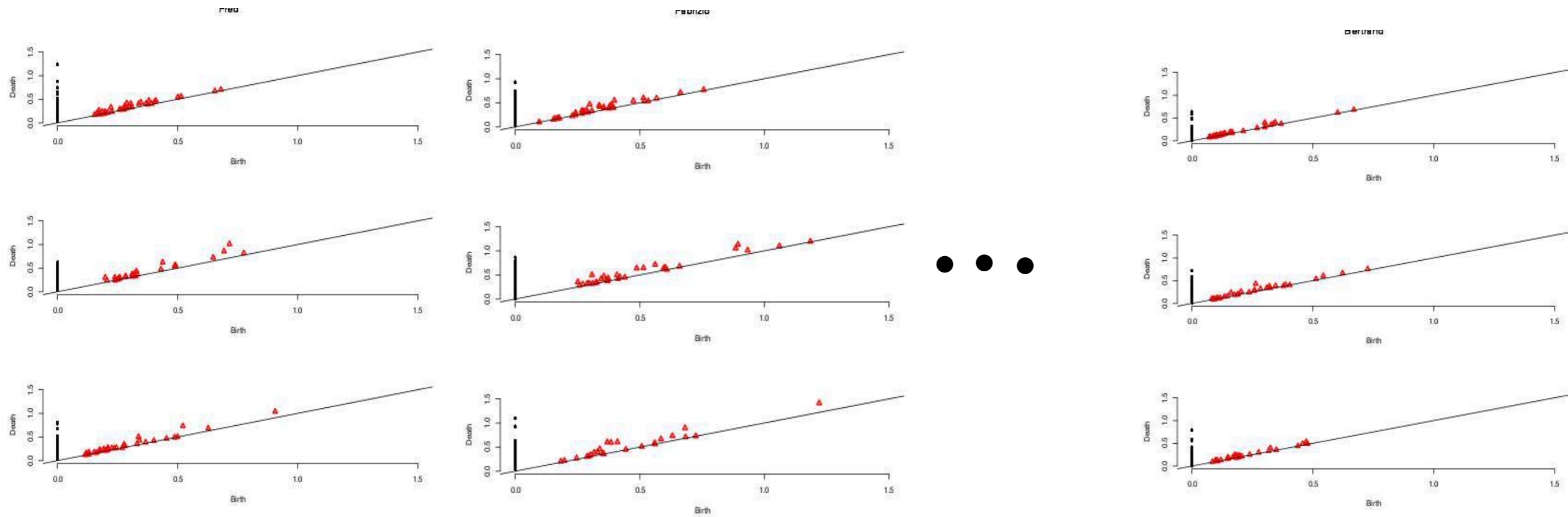
Subsampling method:

- $N$  subsamples  $\mathbb{X}_{b,n}^1, \dots, \mathbb{X}_{b,n}^N$  of size  $b$ .
- Compute  $T_j = d_H(\mathbb{X}_{b,n}^j, \mathbb{X}_n)$ ,  $j = 1, \dots, N$ .
- Compute  $L_b(t) = \frac{1}{N} \sum_{j=1}^N \mathbb{1}_{T_j > t}$ ,
- Take  $c_b = 2L_b^{-1}(\alpha)$ .

If  $P$  satisfies an  $(a, b)$  standard assumption then, for  $n$  large enough :

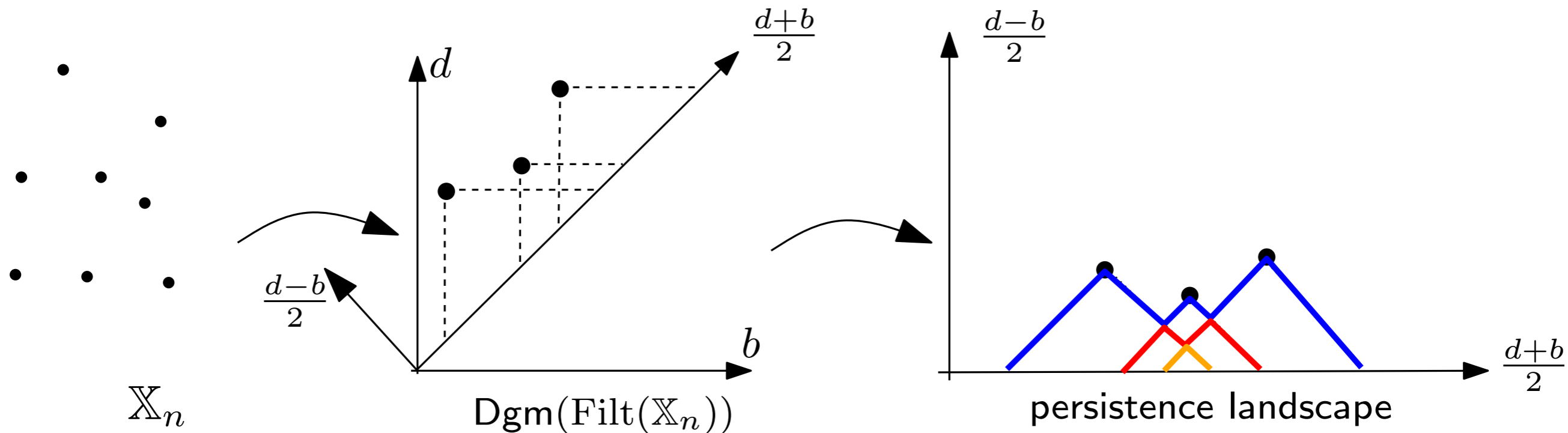
$$\begin{aligned} P(W_\infty(\text{Dgm}(\text{Filt}(\mathbb{X}_\mu)), \text{Dgm}(\text{Filt}(\mathbb{X}_n))) > c_b) &\leq P(d_H(\mathbb{X}_\mu, \mathbb{X}_n) > c_b) \\ &\leq \alpha + O\left(\frac{b}{n}\right)^{1/4} \end{aligned}$$

# Central tendency for persistent homology



- Frechet mean [Turner 2014]
- Use an alternative descriptor of persistence : Persistence landscapes [Bubenik, 2015]

# Persistence landscapes [Bubnik JMLR 2015]



$$\text{Dgm} = \left\{ \left( \frac{d_i + b_i}{2}, \frac{d_i + b_i}{2} \right), i \in I \right\}$$

Persistence landscape  $\lambda$  of Dgm:

$$\lambda(k, t) = \text{kmax}_{p \in D} \Lambda_p(t), \quad t \in \mathbb{R}, k \in \mathbb{N},$$

where  $\text{kmax}$  is  $k$ -th largest value in the set.

For  $p = \left( \frac{b+d}{2}, \frac{d-b}{2} \right) \in \text{Dgm}$ ,

$$\Lambda_p(t) = \begin{cases} t - b & t \in [b, \frac{b+d}{2}] \\ d - t & t \in (\frac{b+d}{2}, d] \\ 0 & \text{otherwise.} \end{cases}$$

**Stability:** For any  $t \in \mathbb{R}$  and any  $k \in \mathbb{N}$ ,  $|\lambda(k, t) - \lambda'(k, t)| \leq d_b(\text{Dgm}, \text{Dgm}')$ .

# Subsampling methods for pers. homology [Chazal ICML 2015]

joint work with F. Chazal, B. Fasy, F. Lecci, A. Rinaldo and L. Wasserman

- Let  $X = \{X_1, \dots, X_m\}$  sampled from  $\mu$ .
- $\lambda_X$ : corresponding persistence landscape.
- $\Psi_\mu^m$ : the measure induced by  $\mu^{\otimes m}$  on the space of persistence landscapes.
- We consider the point-wise expectations of the (random) persistence landscape under this measure:

$$\mathbb{E}_{\Psi_\mu^m} [\lambda_X(t)], t \in [0, T]$$

- For  $S_1^m, \dots, S_\ell^m$  some independent samples of size  $m$  from  $\mu^{\otimes m}$ , the empirical counterpart of  $\mathbb{E}_{\Psi_\mu^m} [\lambda_X(t)]$  is

$$\overline{\lambda_\ell^m}(t) = \frac{1}{\ell} \sum_{i=1}^{\ell} \lambda_{S_i^m}(t), \quad \text{for all } t \in [0, T],$$

# Subsampling methods for pers. homology [Chazal ICML 2015]

**Definition:** The  $p$ -th Wasserstein distance between two measures  $\mu, \nu$  defined on  $(\mathbb{M}, \rho)$  is

$$W_{\rho,p}(\mu, \nu) = \left( \inf_{\Pi} \int_{\mathbb{M} \times \mathbb{M}} [\rho(x, y)]^p d\Pi(x, y) \right)^{\frac{1}{p}},$$

where the infimum is taken over all measures on  $\mathbb{M} \times \mathbb{M}$  with marginals  $\mu$  and  $\nu$ .

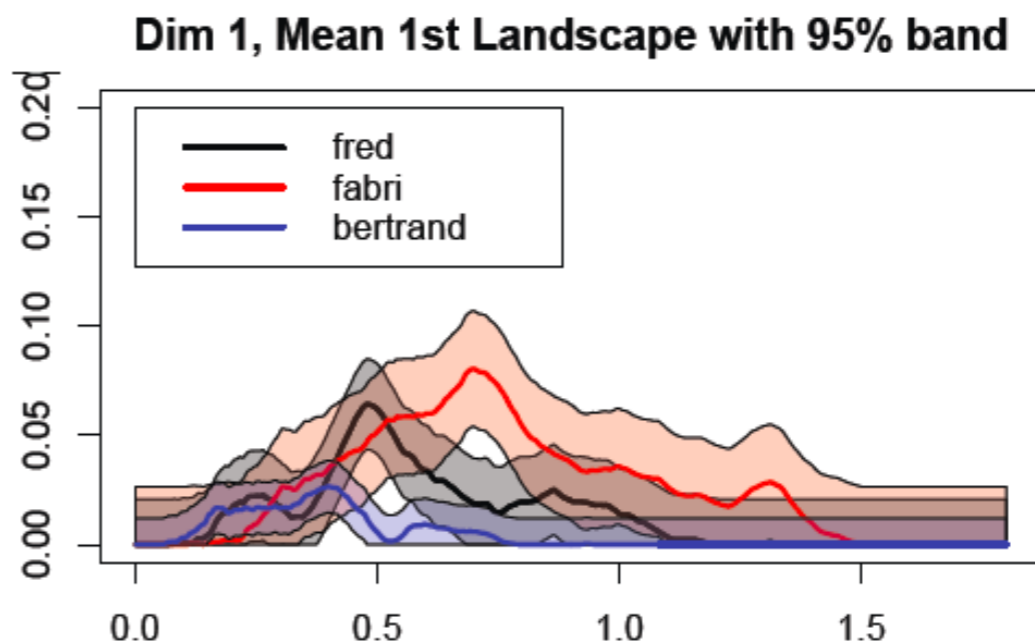
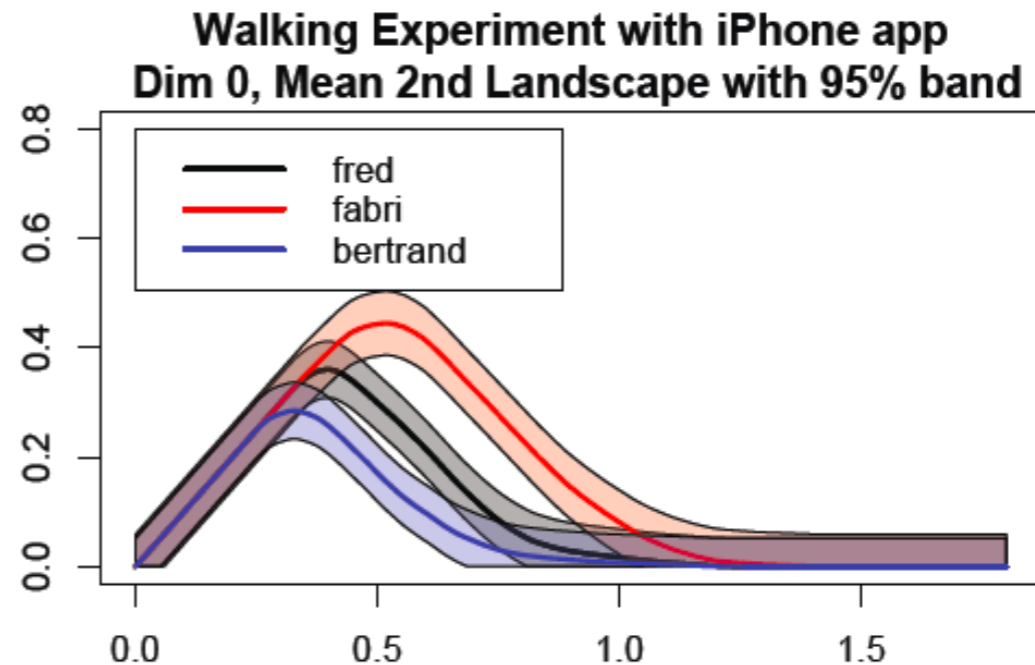
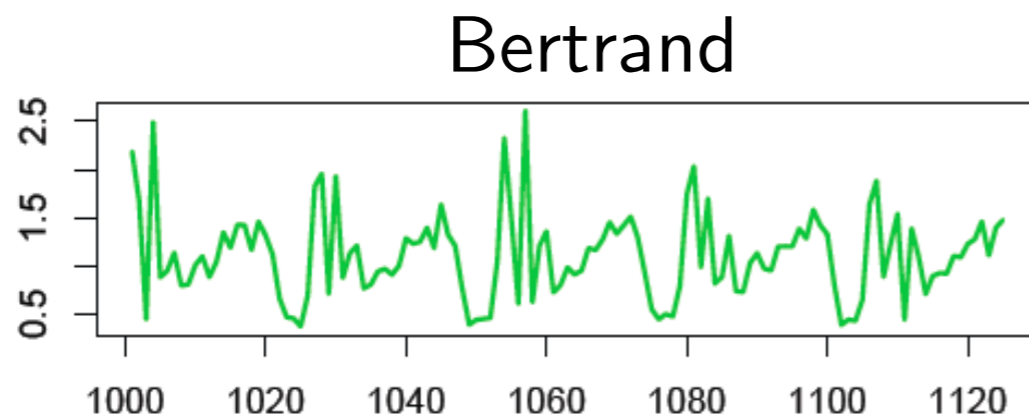
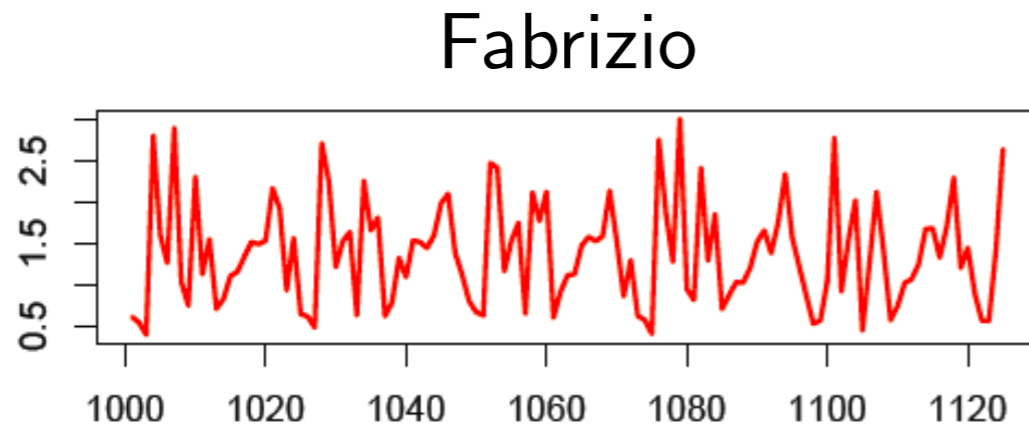
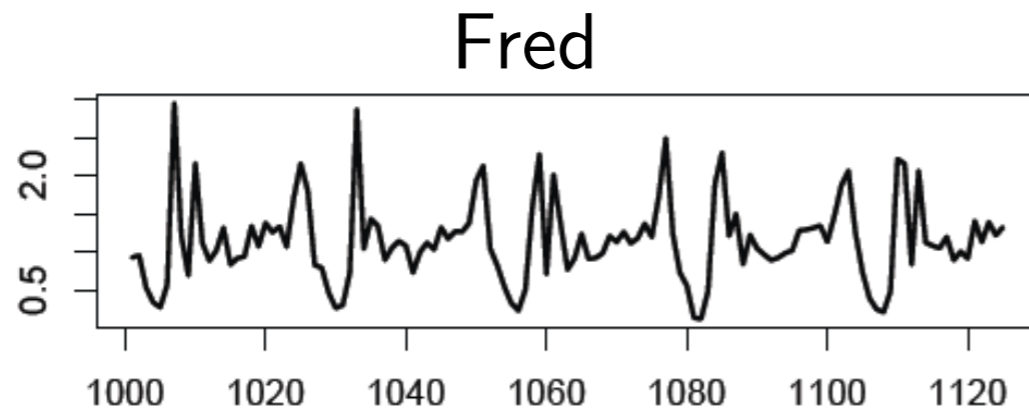
## Stability of the average landscape:

**Theorem:** Let  $X \sim \mu^{\otimes m}$  and  $Y \sim \nu^{\otimes m}$ , where  $\mu$  and  $\nu$  are two probability measures on  $\mathbb{M}$ . For any  $p \geq 1$  we have

$$\left\| \mathbb{E}_{\Psi_{\mu}^m}[\lambda_X] - \mathbb{E}_{\Psi_{\nu}^m}[\lambda_Y] \right\|_{\infty} \leq 2 m^{\frac{1}{p}} W_{\rho,p}(\mu, \nu).$$

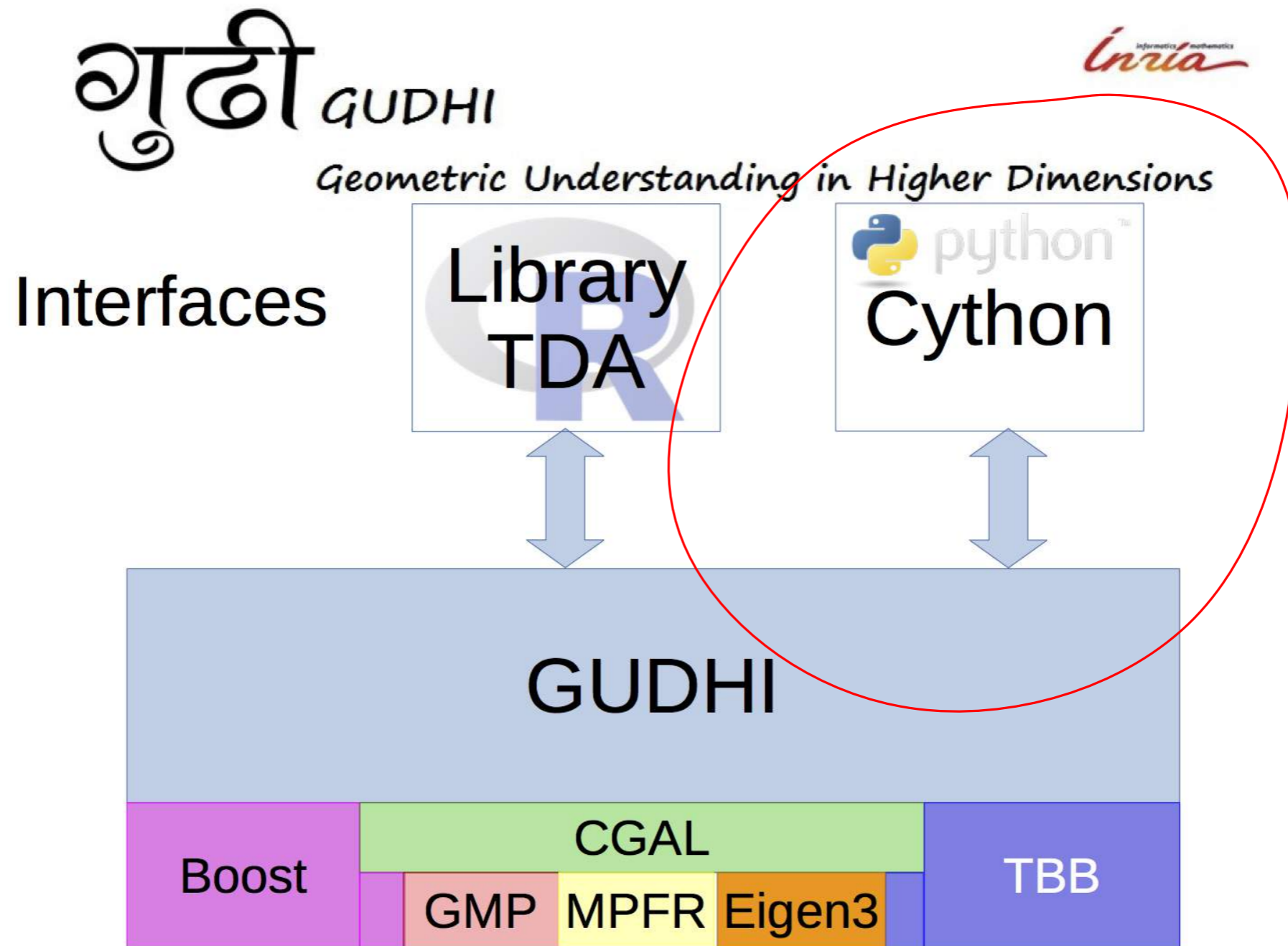
# Subsampling methods for pers. homology [Chazal ICML 2015]

**Application:** Analysis of accelerometer data.



- topological features carry discriminative information
- no registration/calibration preprocessing step needed

# Commercial break: Gudhi with Statistical learning Python Libraries



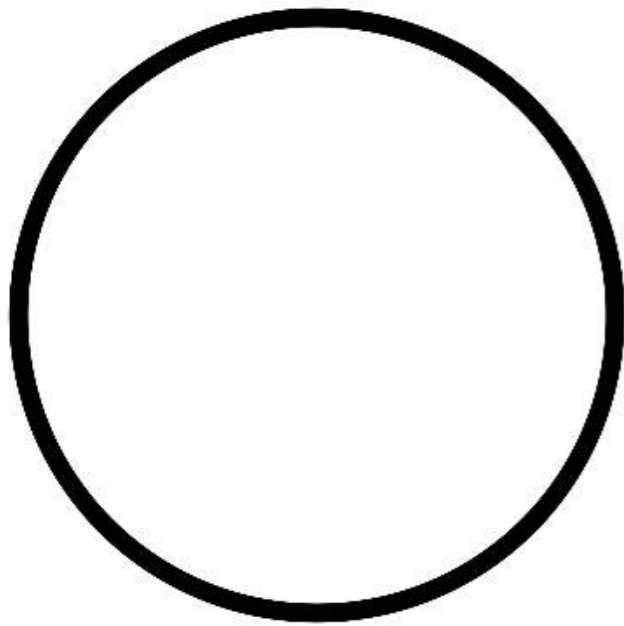
and coming soon : Gudhi Stat with more tools for statistics and TDA.

# Robust TDA

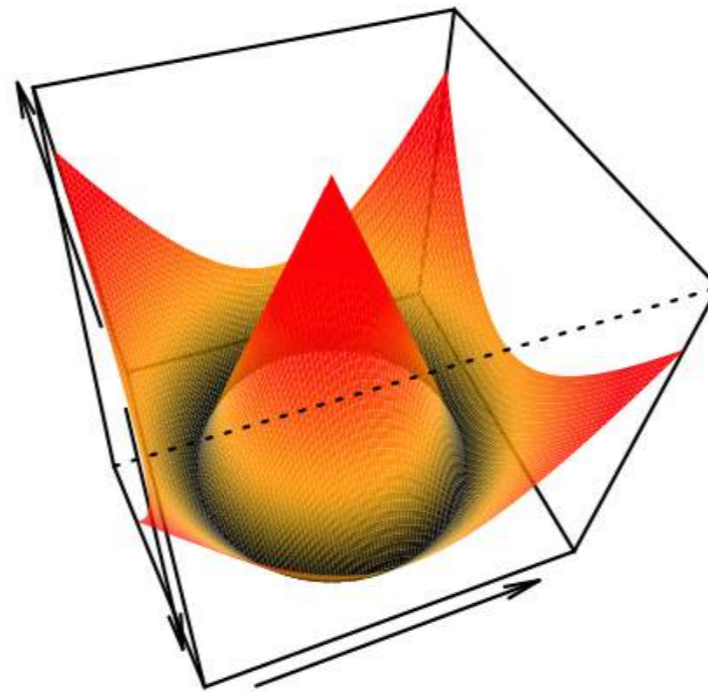


# Standard TDA methods are not robust to outliers

Circle



Distance Function

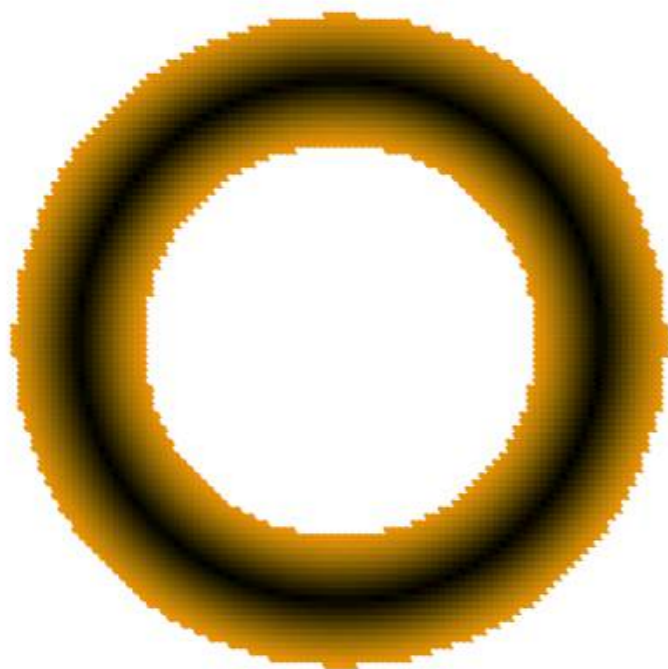


$$\begin{aligned} \mathbb{X}^r &:= \bigcup_{x \in \mathbb{X}} B(x, r) \\ &= d_{\mathbb{X}}^{-1}([0, r]) \end{aligned}$$

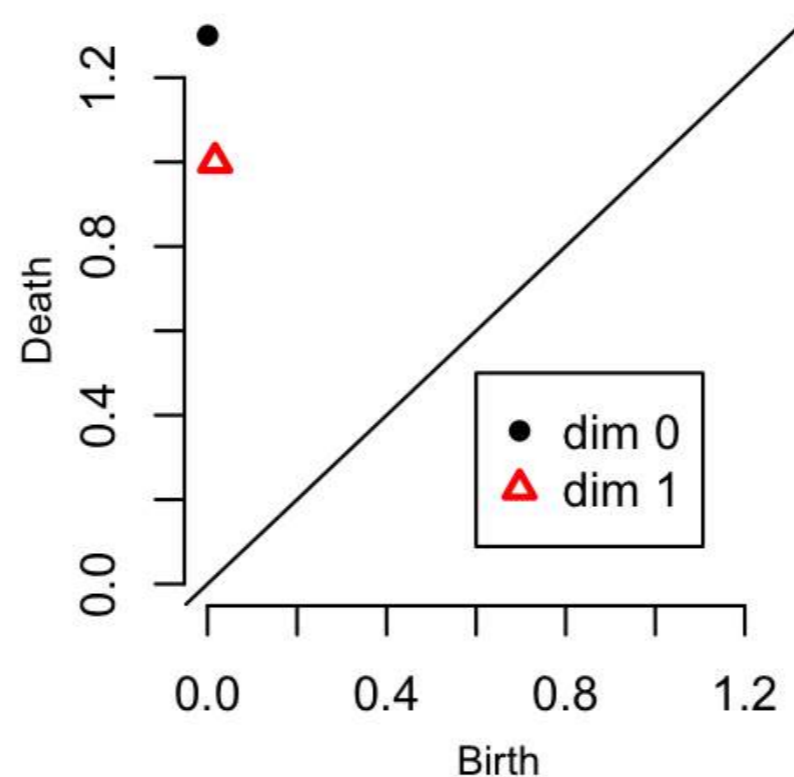
where the distance function  $d_{\mathbb{X}}$  to  $\mathbb{X}$  is

$$d_{\mathbb{X}}(y) = \inf_{x \in \mathbb{X}} \|x - y\|$$

Sublevel Set,  $t=0.25$

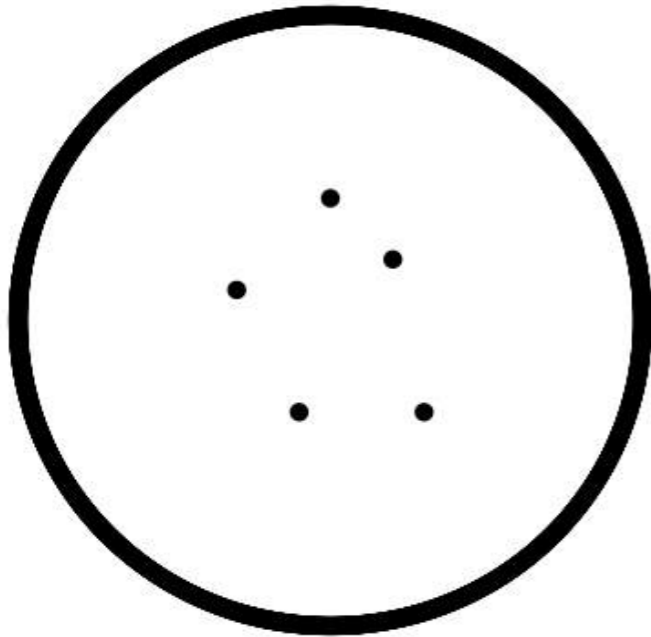


Persistence Diagram

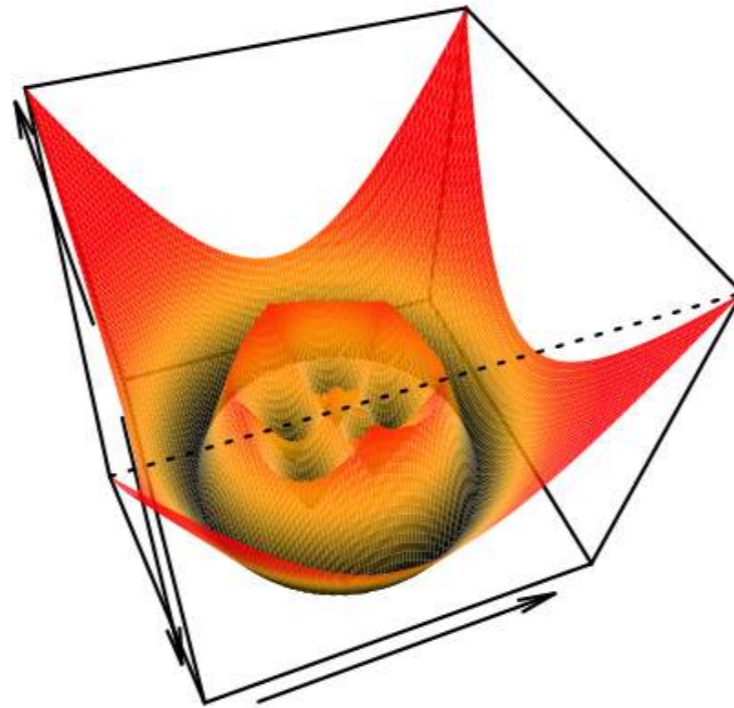


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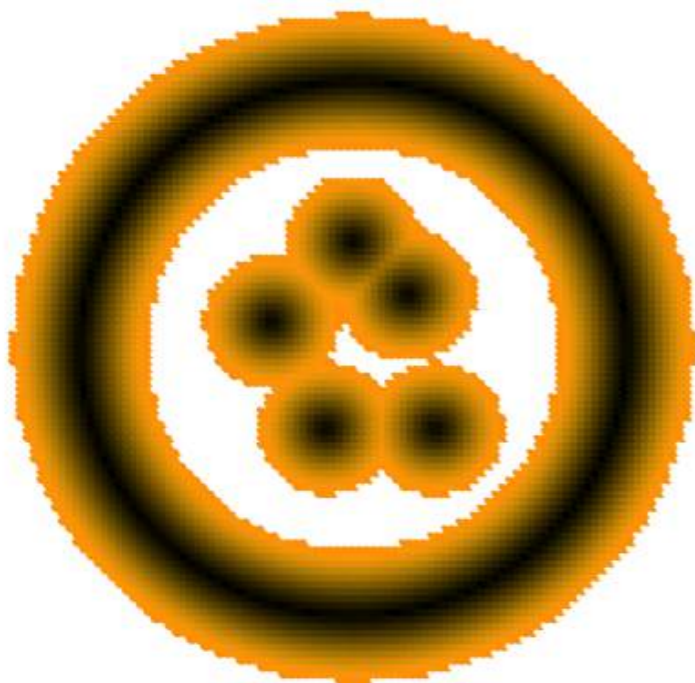
Circle with Outliers



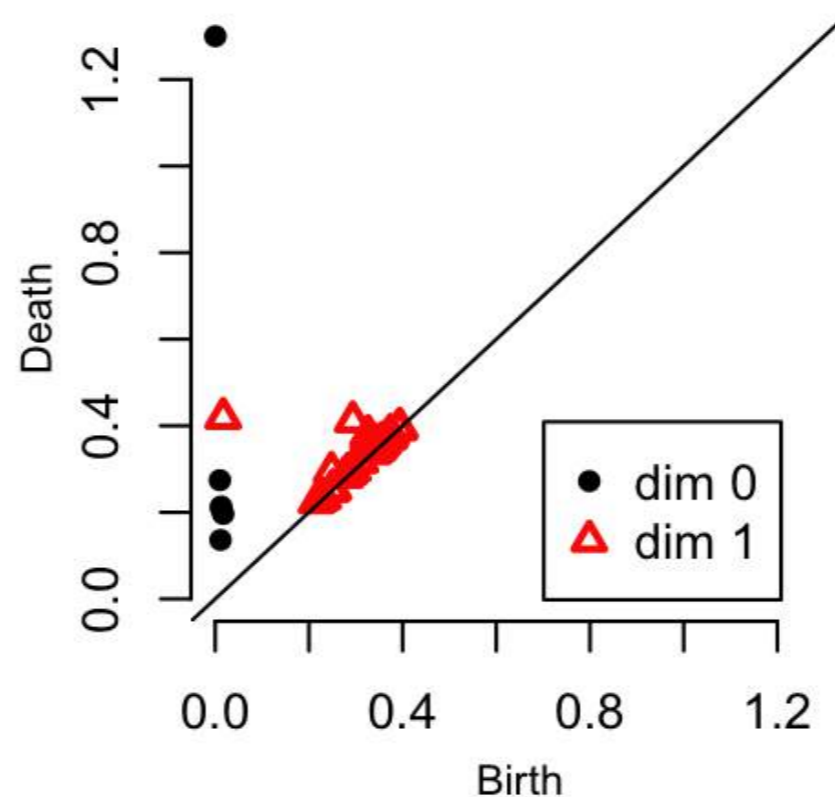
Distance Function



Sublevel Set,  $t=0.25$



Persistence Diagram



$$\begin{aligned} \mathbb{X}^r &:= \bigcup_{x \in \mathbb{X}} B(x, r) \\ &= d_{\mathbb{X}}^{-1}([0, r]) \end{aligned}$$

where the distance function  $d_{\mathbb{X}}$  to  $\mathbb{X}$  is

$$d_{\mathbb{X}}(y) = \inf_{x \in \mathbb{X}} \|x - y\|$$

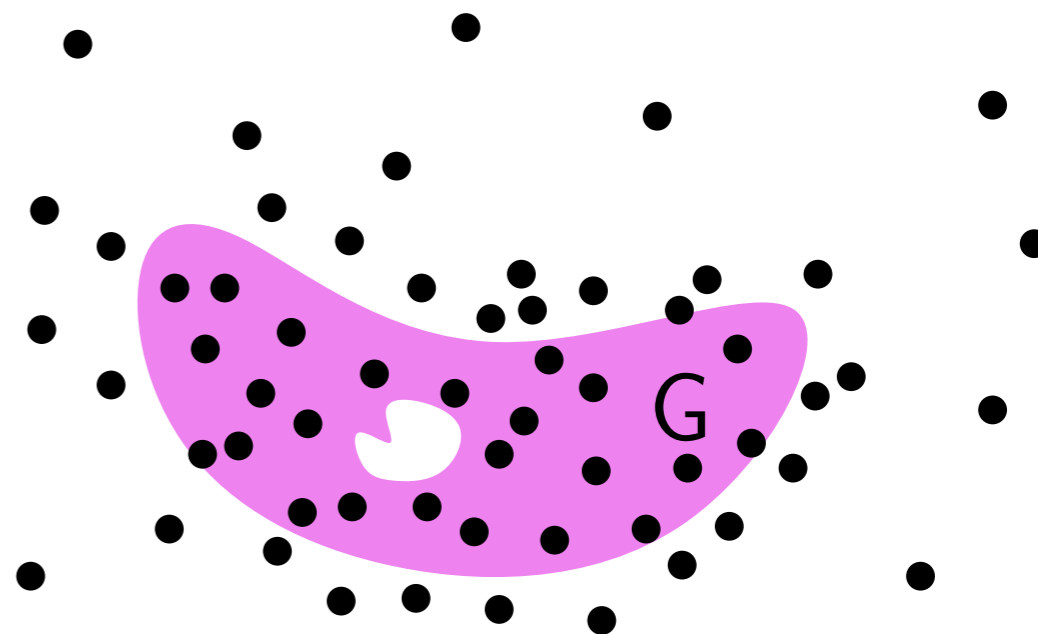
# Some possible “noise models” for geometry

- Additive noise model

$$P = \mu \star \Phi$$

distribution with support  
the geometric shape  $G$

noise distribution

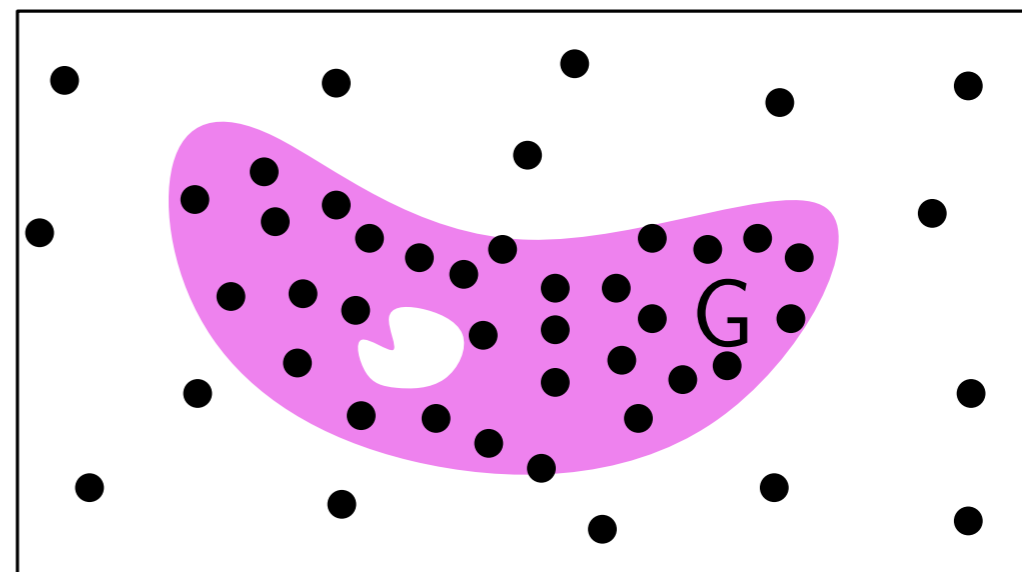


- Clutter noise model

$$P = \pi\mu + (1 - \pi)U$$

distribution with support  
the geometric shape  $G$

Uniform distribu-  
tion on the box

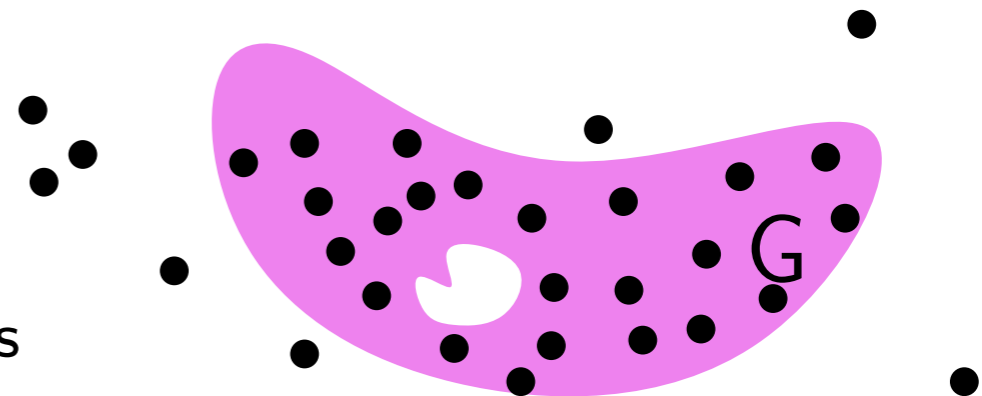


- A few outliers

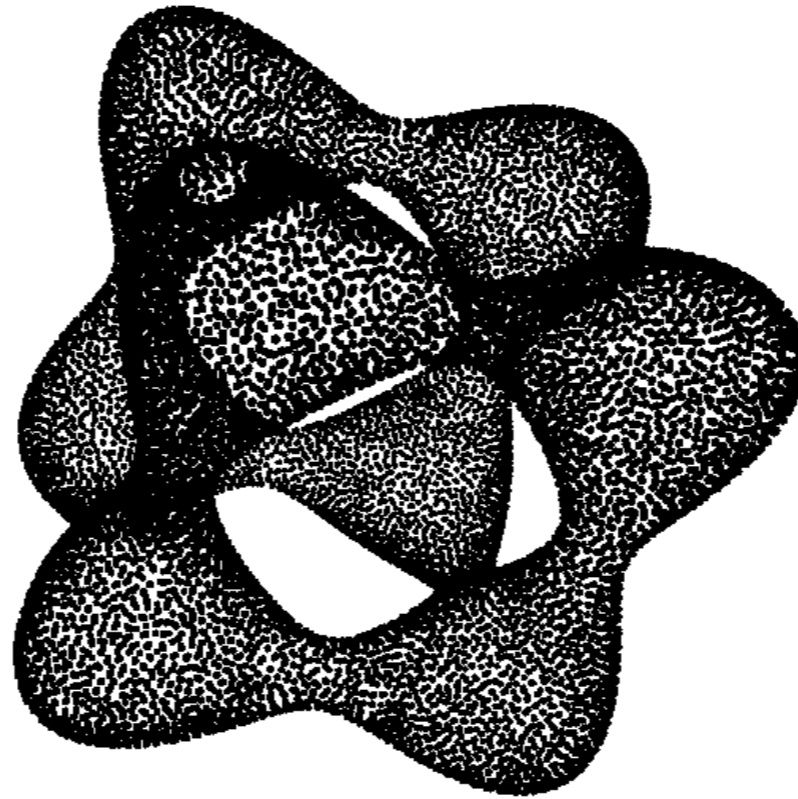
$$P = \pi\mu + (1 - \pi)\Psi$$

distribution with support  
the geometric shape  $G$

Distribution of outliers



# Robust TDA with an alternative distance function ?



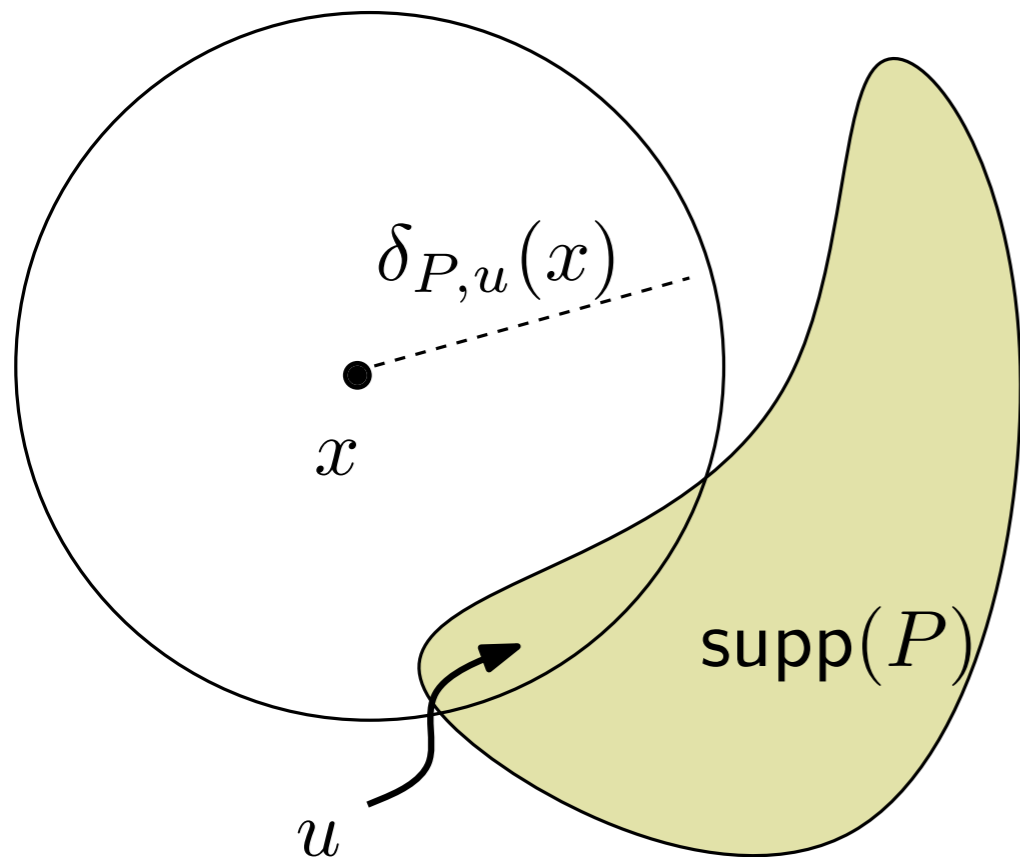
We would like to consider the sub levels of an alternative distance function related to the sampling measure, which support is  $\mathbb{X}$ , or close to  $\mathbb{X}$ .

# Distance To Measure [Chazal 11 FoCM]

## Preliminary distance function to a measure $P$ :

Let  $u \in ]0, 1[$  be a positive mass, and  $P$  a probability measure on  $\mathbb{R}^d$ :

$$\delta_{P,u}(x) = \inf \{r > 0 : P(B(x, r)) \geq u\}$$



$\delta_{P,u}$  is the smallest distance needed to capture a mass of at least  $u$ .

$\delta_{P,u}$  is the quantile function at  $u$  of the r.v.

$$\|x - X\|$$

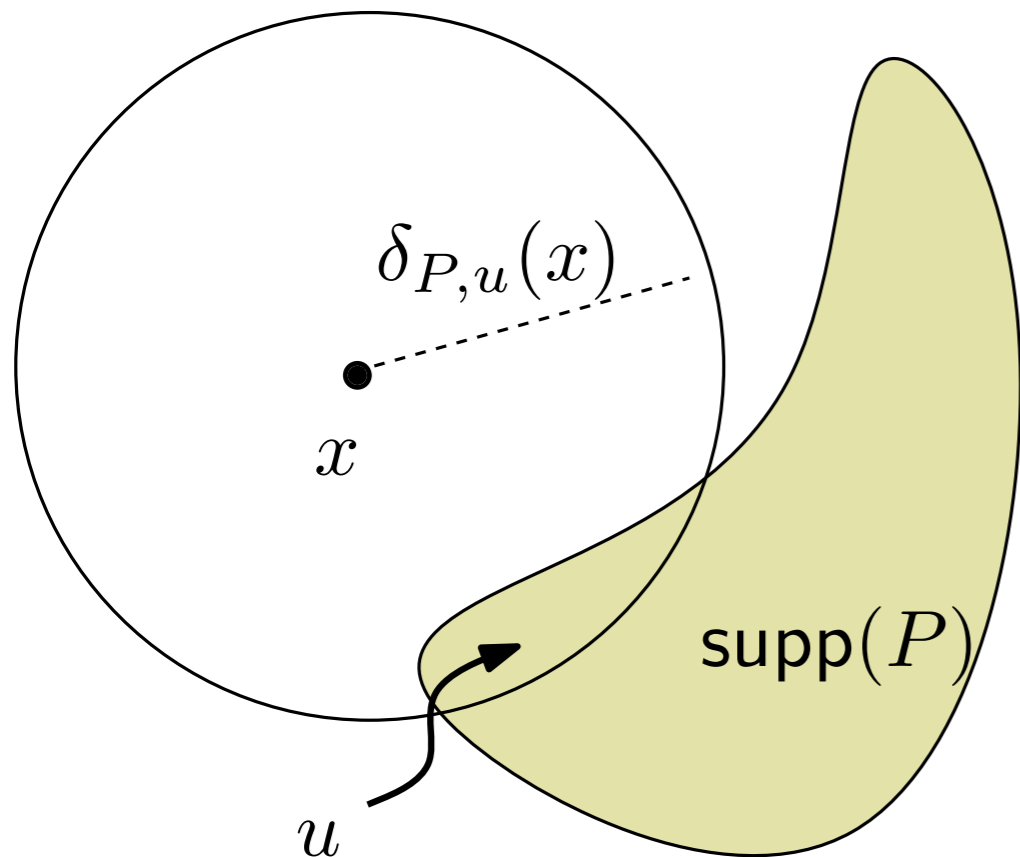
where  $X \sim P$ .

# Distance To Measure [Chazal 11 FoCM]

## Preliminary distance function to a measure $P$ :

Let  $u \in ]0, 1[$  be a positive mass, and  $P$  a probability measure on  $\mathbb{R}^d$ :

$$\delta_{P,u}(x) = \inf \{r > 0 : P(B(x,r)) \geq u\}$$



**Definition:** Given a probability measure  $P$  on  $\mathbb{R}^d$  and  $m > 0$ , the distance function to the measure  $P$  (DTM) is defined by

$$d_{P,m} : x \in \mathbb{R}^d \mapsto \left( \frac{1}{m} \int_0^m \delta_{P,u}^2(x) du \right)^{1/2}$$

# Distance To Measure [Chazal 11 FoCM]

## Properties of the DTM :

- Stability under Wassertein perturbations:

$$\|d_{P,m} - d_{Q,m}\|_{\infty} \leq \frac{1}{\sqrt{m}} W_2(P, Q)$$

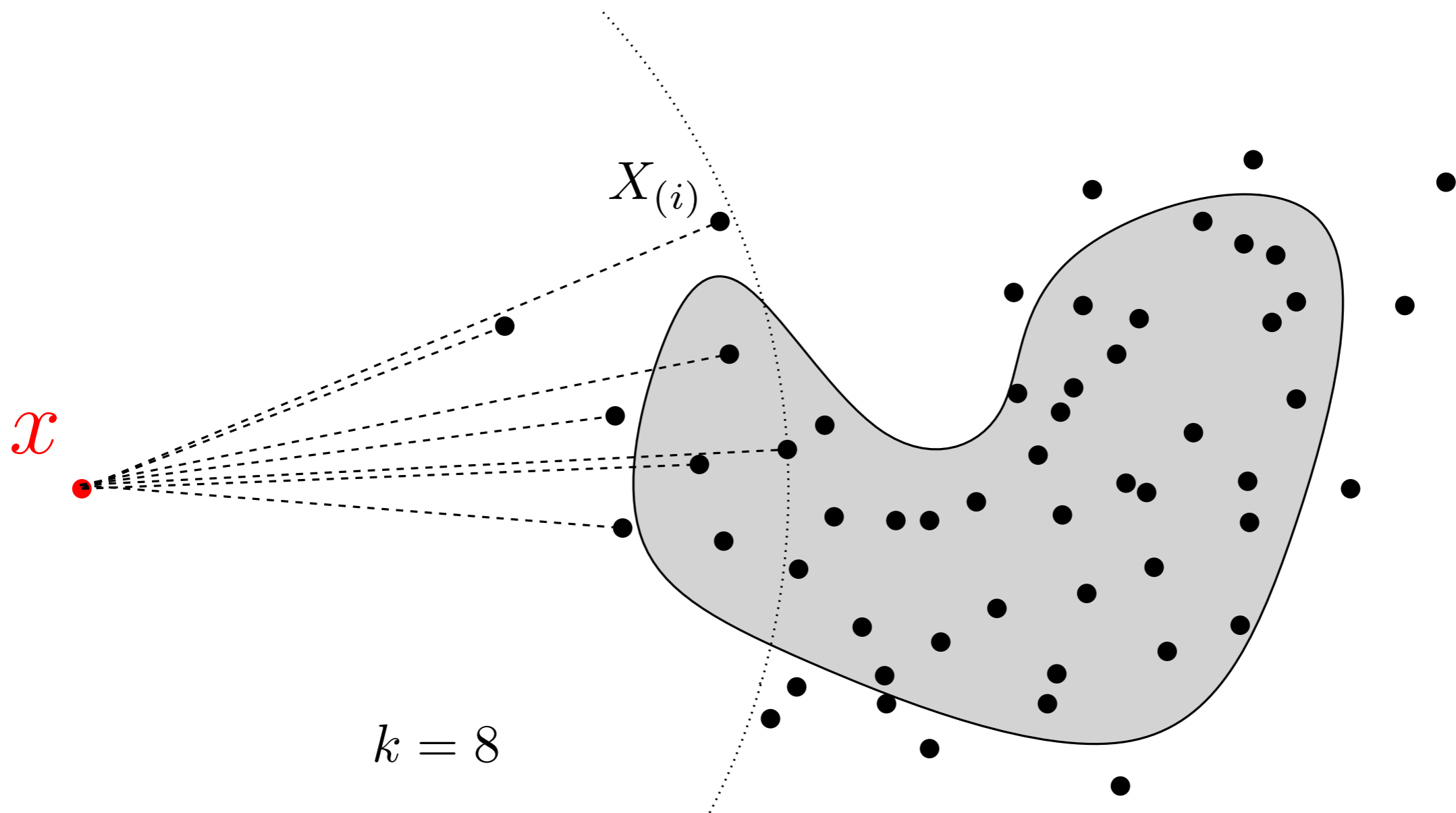
- The function  $x \mapsto d_{P,m}^2(x)$  is semiconcave, this is ensuring strong regularity properties on the geometry of its sublevel sets.
- Consequently, if  $\tilde{P}$  is a probability distribution close to  $P$  for Wasserstein distance  $W_2$ , then the sublevel sets of  $d_{\tilde{P},m}$  provide a topologically correct approximation of the support of  $P$ .

# Distance to The Empirical Measure (DTEM)

Let  $X_1, \dots, X_n$  sample according to  $P$  and let  $P_n$  be the empirical measure.  
Then

$$d_{P_n, \frac{k}{n}}^2(x) = \frac{1}{k} \sum_{i=1}^k \|x - X_{(i)}\|^2$$

where  $\|X_{(1)} - x\| \geq \|X_{(2)} - x\| \geq \dots \geq \|X_{(k)} - x\| \dots \geq \|X_{(n)} - x\|$





# Geometric inference with the DTM

**Theorem:** [Chazal et al., 2011]

Let  $\mu$  be a measure that has dimension at most  $k > 0$  with compact support  $G$  such that  $\text{reach}_\alpha(G) \geq R > 0$  for some  $\alpha > 0$ .

Let  $\nu$  be another measure and  $\varepsilon$  be an upper bound on the uniform distance between  $d_G$  and  $d_{\nu, m_0}$ . Then, for any  $r \in [4\varepsilon/\alpha^2, R - 3\varepsilon]$  and any  $\eta \in ]0, R[$ , the  $r$ -sublevel sets of  $d_{\mu, m_0}$  and the  $\eta$ -sublevel sets of  $d_G$  are homotopy equivalent as soon as:

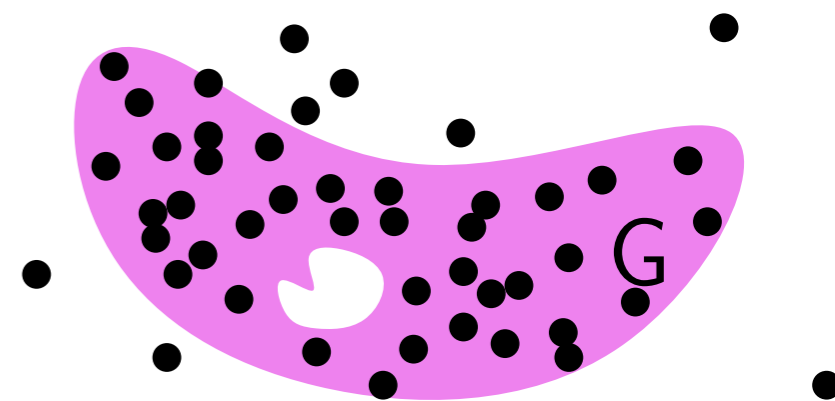
$$W_2(\mu, \nu) \leq \frac{R\sqrt{m_0}}{5 + 4/\alpha^2} - C(\mu)^{-1/k} m_0^{1/k+1/2}.$$

In practice :  $X_1 \dots X_n$  sampled according to  $P$ .

Assume  $W_2(P, \mu)$  small.

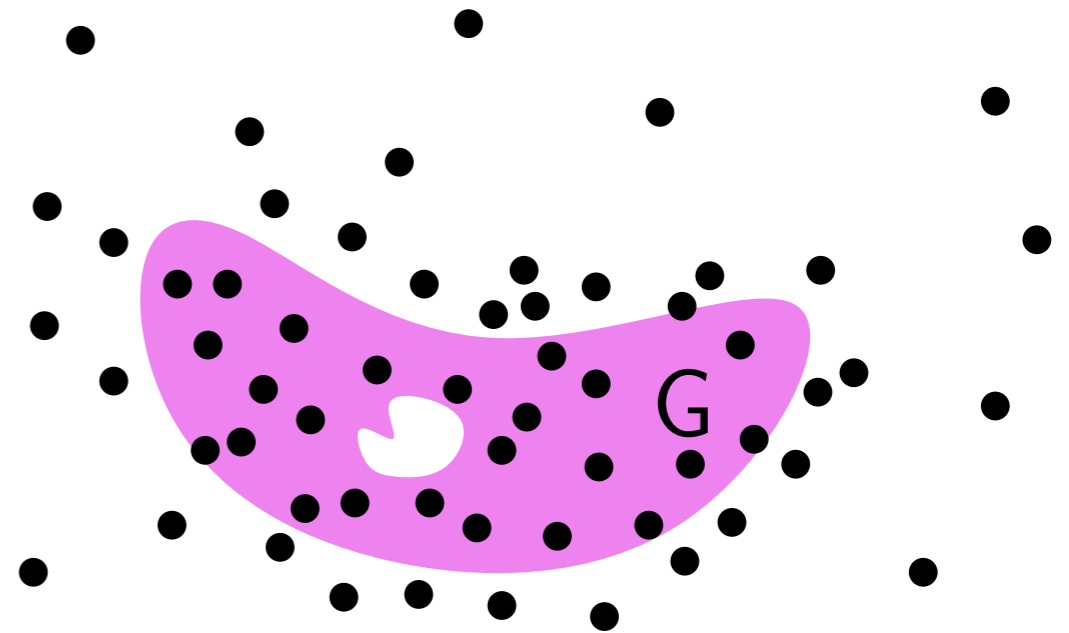
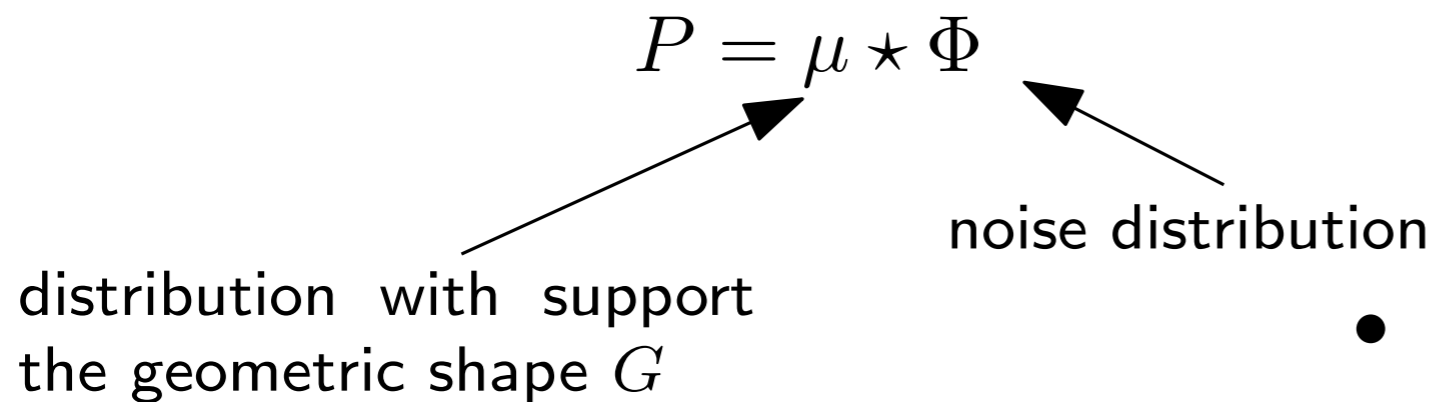
$P_n = \sum_{i=1}^n \delta_{X_i}$  : empirical measure.

Then for  $n$  large enough,  $W_2(P_n, \mu)$  is small and the sublevel sets of  $d_{P_n, m}$  provide a topologically correct approximation of  $G$ .



# Wasserstein deconvolution and DTM denoising

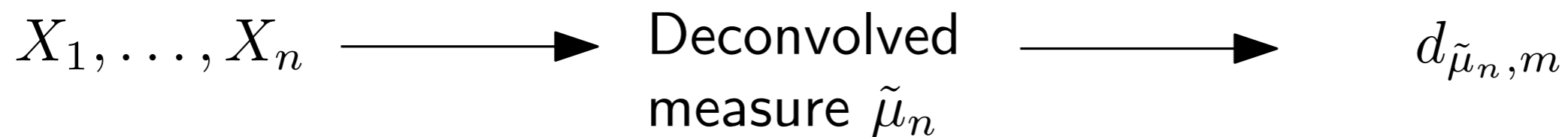
Additive noise model



In this case,  $W_2(P, \mu)$  can be large.

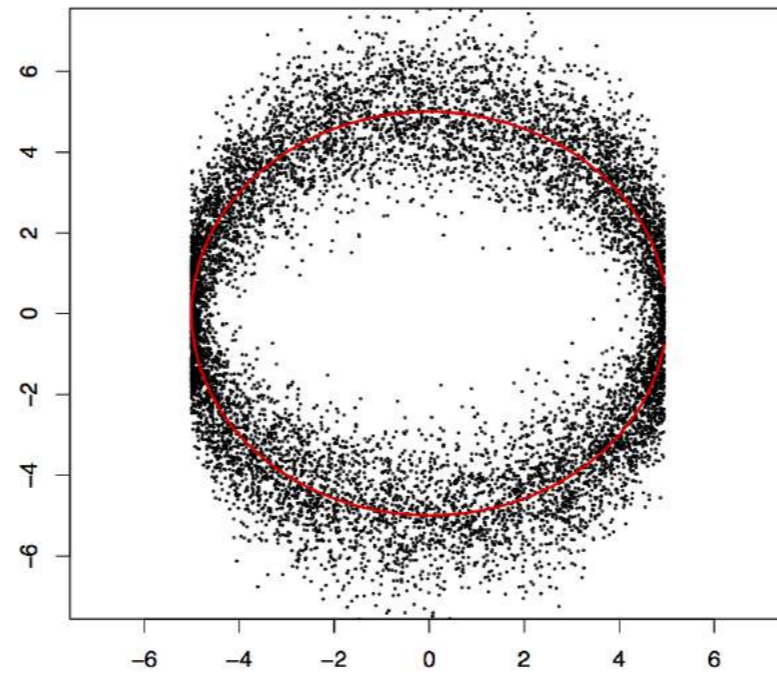
Ideally we would like to denoise directly  $d_{P_n, n}$ , but this can be hardly achieved because the DTM is not a linear functional of the measure.

Alternative approach : deconvolve the observed measure [Caillerie EJS 2011]

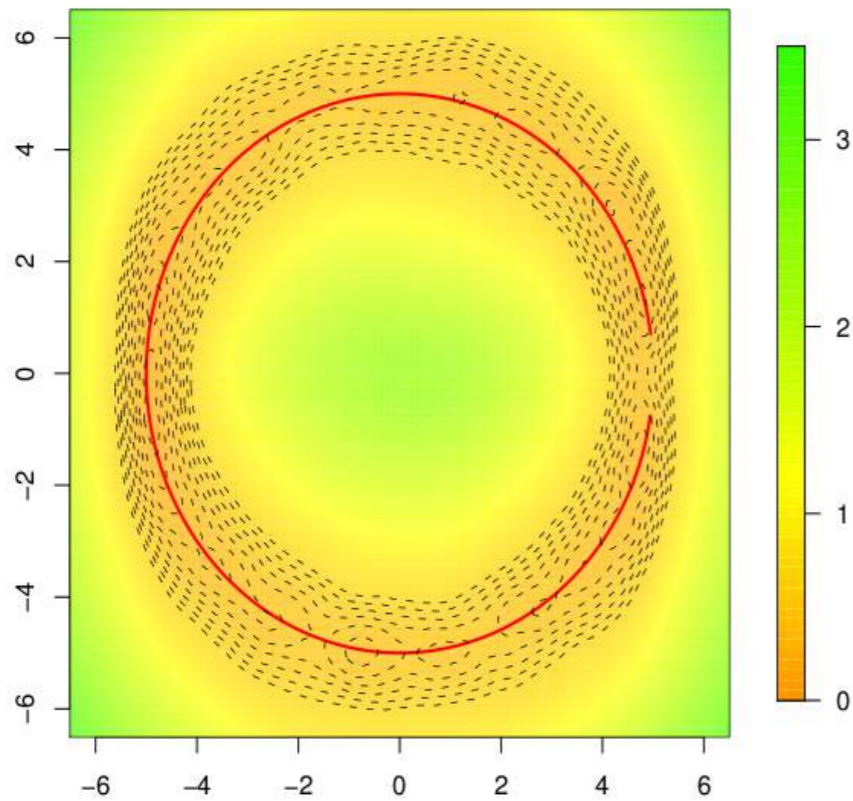


$$\begin{array}{ccc}
 W_2(P_n, \mu) \geq W_2(\tilde{\mu}_n, \mu) \geq \sqrt{m} \|d_{\tilde{\mu}_n, m} - d_\mu\|_\infty \\
 (\not\rightarrow 0) & (\rightarrow 0) & 
 \end{array}$$

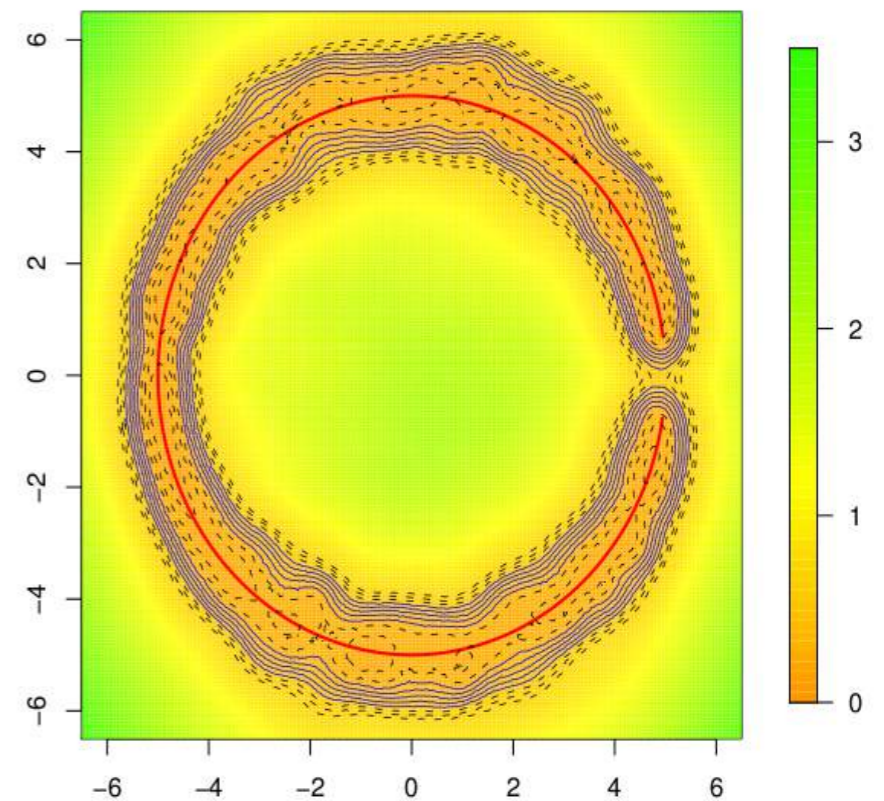
# Wasserstein deconvolution and DTM denoising



Distance to the empirical measure

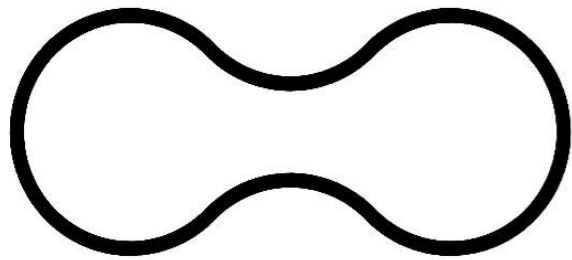


Distance to the estimator

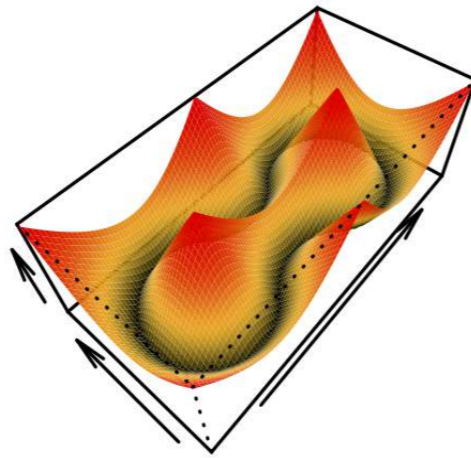


# DTM and persistent homology

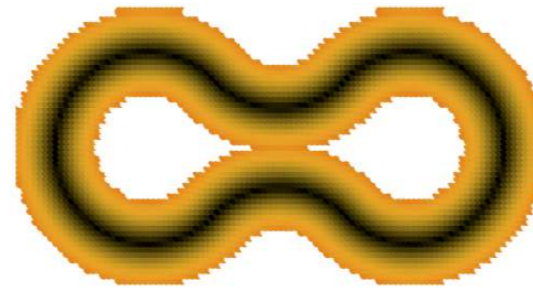
Cassini Curve



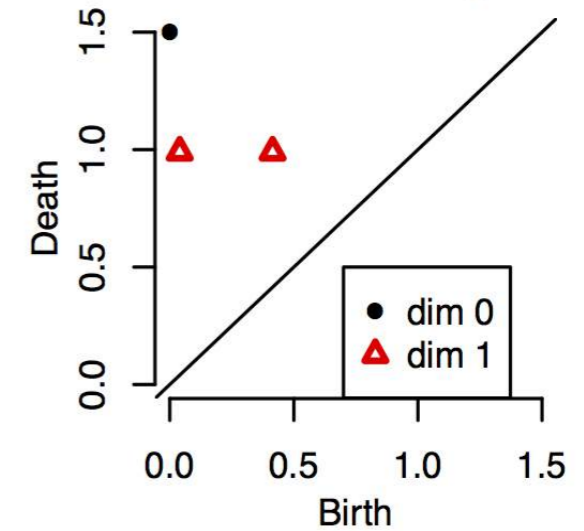
Distance Function



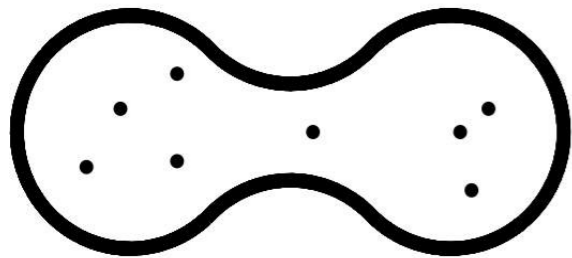
Sublevel Set,  $t=0.45$



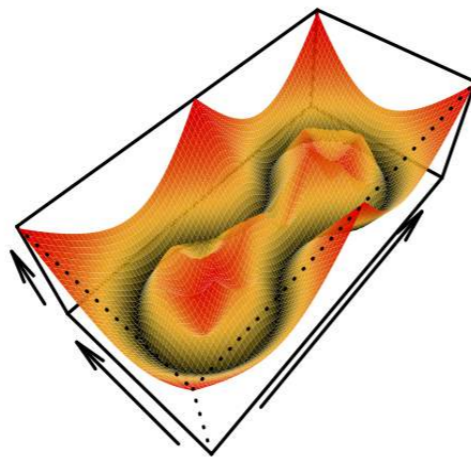
Persistence Diagram



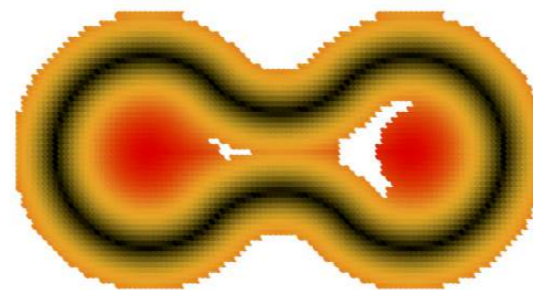
Cassini with Outliers



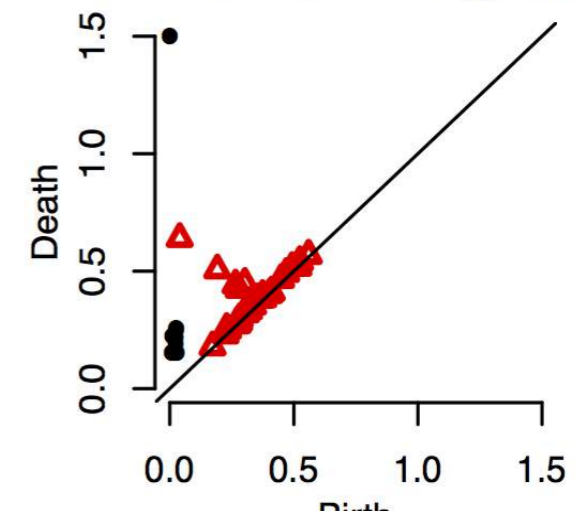
Distance Function



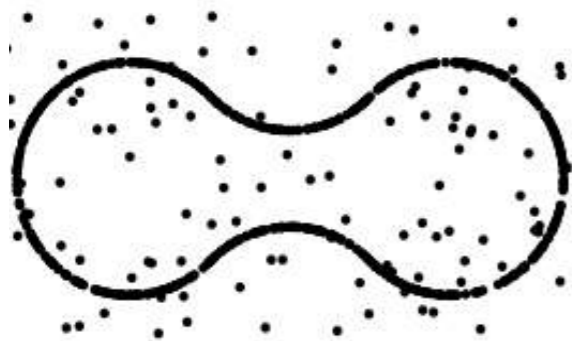
Sublevel Set,  $t=0.45$



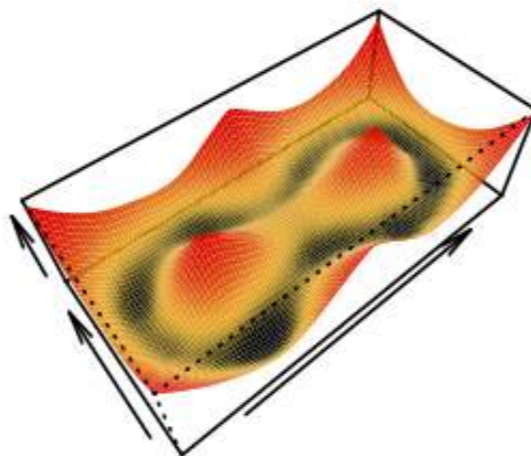
Persistence Diagram



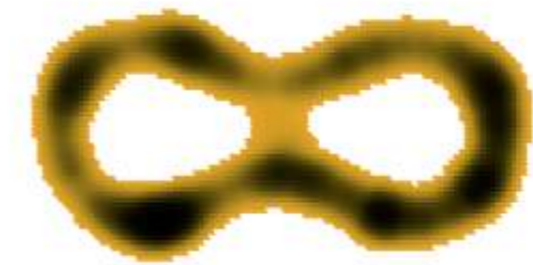
Cassini with Noise



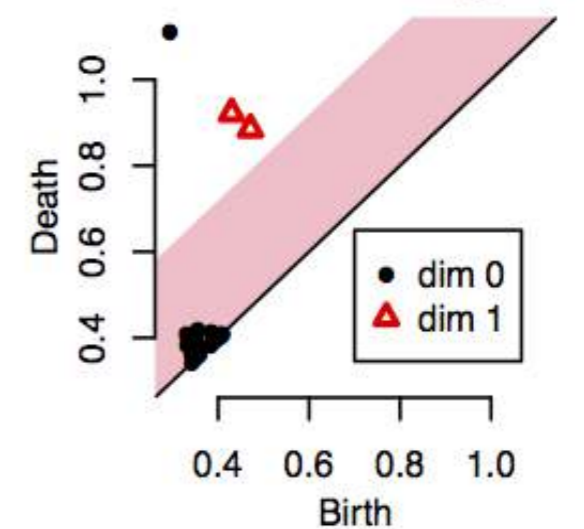
DTM



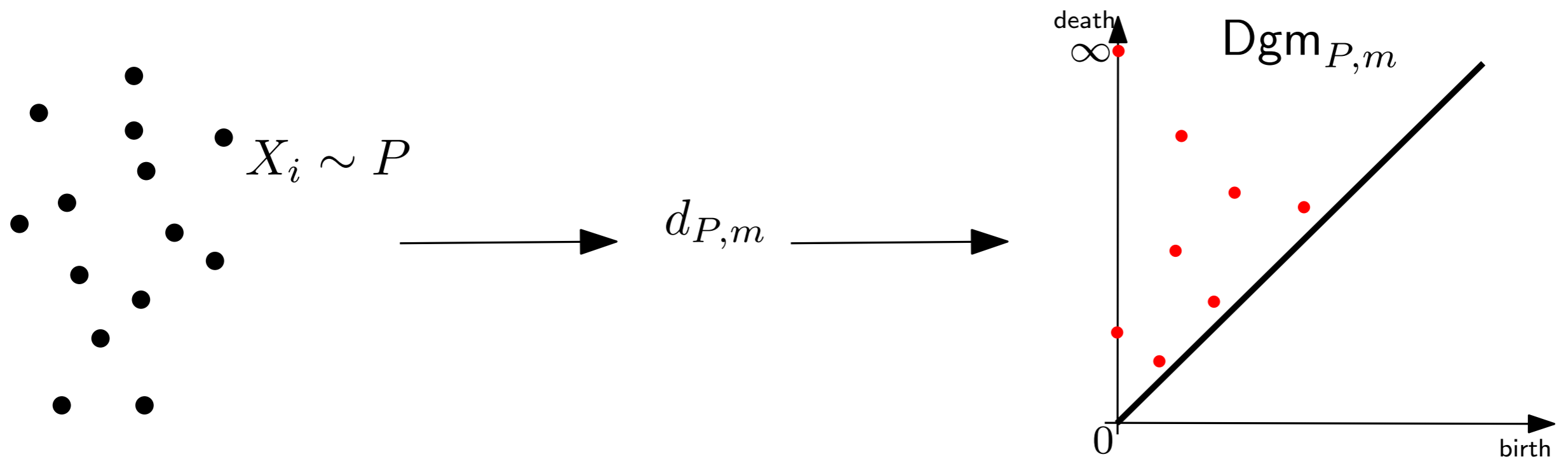
Sublevel Set,  $t=0.5$



Persistence Diagram



# DTM and persistent homology



$$d_b (Dgm_{P,m}, Dgm_{Q,m}) \leq \|d_{P,m} - d_{Q,m}\|_\infty \leq \frac{1}{\sqrt{m}} W_2(P, Q)$$

Take  $Q = P_n \dots$

Wasserstein Stability of the DTM  
[Chazal et al., 2012]

Stability of Persistent homology [Cohen-Steiner et al., 2005, Chazal et al., 2012]

# Estimation of the DTM via the empirical DTM

[Chazal EJS 17, Chazal JMLR 17]

Quantity of interest:

$$d_{P_n, \frac{k}{n}}^2(x) - d_{P, \frac{k}{n}}^2(x)$$

- Observe that

$$d_{P, m}^2(x) = \frac{1}{m} \int_0^m F_x^{-1}(u) du$$

where  $F_x$  is the cdf of  $\|x - X\|^2$  with  $X \sim P$ .

- The distance to the empirical measure is the empirical counter part of the distance to  $P$ :

$$d_{P_n, m}^2(x) = \frac{1}{m} \int_0^m F_{x, n}^{-1}(u) du$$

where  $F_{x, n}$  is the cdf of  $\|x - X\|^2$  with  $X \sim P_n$ .

- Finally we get that

$$d_{P_n, \frac{k}{n}}^2(x) - d_{P, \frac{k}{n}}^2(x) = \frac{1}{m} \int_0^m \{F_{x, n}^{-1}(u) - F_x^{-1}(u)\} du$$

# Estimation of the DTM via the empirical DTM

[Chazal EJS 17, Chazal JMLR 17]

Quantity of interest:

$$d_{P_{n, \frac{k}{n}}}^2(x) - d_{P, \frac{k}{n}}^2(x)$$

Two complementary approaches of the problem:

- Asymptotic approach :  $\frac{k_n}{n} = m$  is fixed and  $n$  tends to infinity.
- Non asymptotic approach :  $n$  is fixed, and we want a tight control over the fluctuations of the empirical DTM, in function of  $k$ , which can be taken very small.

We **do not use Wasserstein stability** for either of the two approaches. Wasserstein rates of convergence [Fournier and Guillin, 2013 ; Dereich et al., 2013] do not provide tight rates for the DTM in this context.

# Bootstrap and significance of topological features

**Aim** : studying the persistent homology of the sub-levels of the DTM and providing confidence regions.

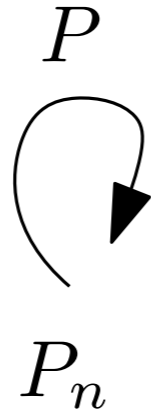
Two alternative bootstrap methods :

- by bootstrapping the DTM
- Bottleneck Bootstrap



# Bootstrap and significance of topological features

## Bootstrapping the DTM



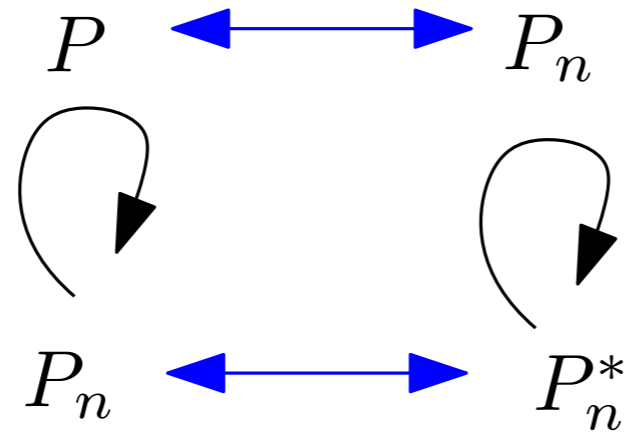
# Bootstrap and significance of topological features

## Bootstrapping the DTM



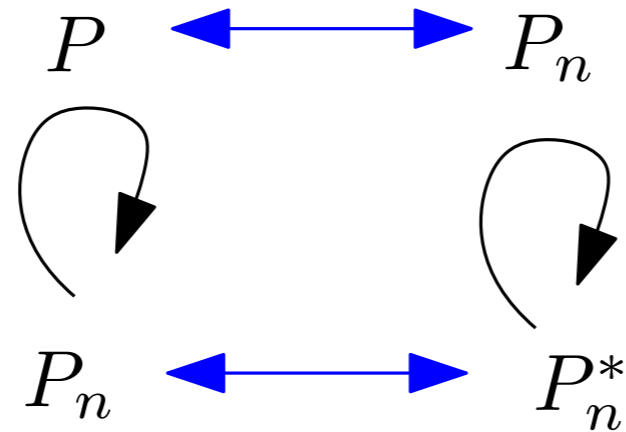
# Bootstrap and significance of topological features

## Bootstrapping the DTM



# Bootstrap and significance of topological features

## Bootstrapping the DTM



$$\Phi(P) \longleftrightarrow \Phi(P_n)$$

$$\Phi(P_n) \longleftrightarrow \Phi(P_n^*)$$

# Bootstrap and significance of topological features

## Bootstrapping the DTM

For  $m \in (0, 1)$ , define  $c_\alpha$  by

$$\mathbb{P} \left( \sqrt{n} \|d_{P,m}^2 - d_{P_n,m}^2\|_\infty > c_\alpha \right) = \alpha.$$

Let  $X_1^*, \dots, X_n^*$  be a sample from  $P_n$ , and let  $P_n^*$  be the corresponding (bootstrap) empirical measure.

We consider the bootstrap quantity  $d_{P_n^*,m}(x)$  of  $d_{P_n,m}$ .

The bootstrap estimate  $\hat{c}_\alpha$  is defined by

$$\mathbb{P} \left( \sqrt{n} \|d_{P_n,m}^2 - d_{P_n^*,m}^2\|_\infty > \hat{c}_\alpha \mid X_1, \dots, X_n \right) = \alpha$$

where  $\hat{c}_\alpha$  can be approximated by Monte Carlo.

**Theorem:** If  $F_x^{-1}$  is regular enough, the DTM is Hadamard differentiable at  $P$ . Consequently, the bootstrap method for the DTM is asymptotically valid.

# Bootstrap and significance of topological features

## Bootstrapping the DTM

$D_{\text{gm}}$  : persistence diagram of the sub-levels of  $d_{P,m}$

$\widehat{D}_{\text{gm}}$  : persistence diagram of the sub-levels of  $d_{P_n,m}$ .

Let

$$\mathcal{C}_n = \left\{ E \in \mathcal{D}\text{diag} : d_b(\widehat{D}_{\text{gm}}, E) \leq \frac{\hat{c}_\alpha}{\sqrt{n}} \right\},$$

where  $\mathcal{D}\text{diag}$  is the set of all the persistence diagrams.

Then,

Bootstrap estimate

$$\mathbb{P}(D_{\text{gm}} \in \mathcal{C}_n) = \mathbb{P} \left( d_b(D_{\text{gm}}, \widehat{D}_{\text{gm}}) \leq \frac{\hat{c}_\alpha}{\sqrt{n}} \right) \geq \mathbb{P} \left( \|d_{P,m}^2 - d_{P_n,m}^2\|_\infty \leq \frac{\hat{c}_\alpha}{\sqrt{n}} \right)$$

# Bootstrap and significance of topological features

## The Bottleneck Bootstrap

$D_{\text{gsm}}$  : persistence diagram of the sub-levels of  $d_{P,m}$

$\widehat{D_{\text{gsm}}}$  : persistence diagram of the sub-levels of  $d_{P_n,m}$ .

$\widehat{D_{\text{gsm}}}^*$  : persistence diagram of the sub-levels of  $d_{P_n^*,m}$ .

We directly bootstrap in the set of the persistence diagram by considering the random quantity  $d_b(\widehat{D_{\text{gsm}}}^*, \widehat{D_{\text{gsm}}})$ . We define  $\hat{t}_\alpha$  by

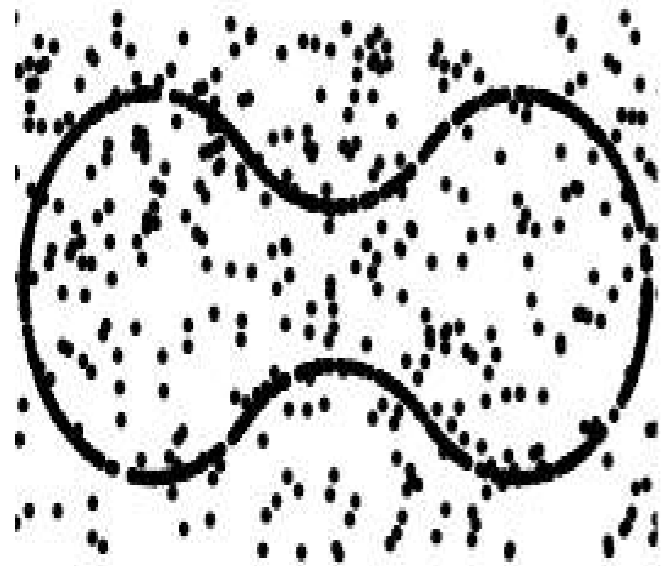
$$\mathbb{P} \left( \sqrt{n} d_b(\widehat{D_{\text{gsm}}}^*, \widehat{D_{\text{gsm}}}) > \hat{t}_\alpha \mid X_1, \dots, X_n \right) = \alpha.$$

The quantile  $\hat{t}_\alpha$  can be estimated by Monte Carlo.

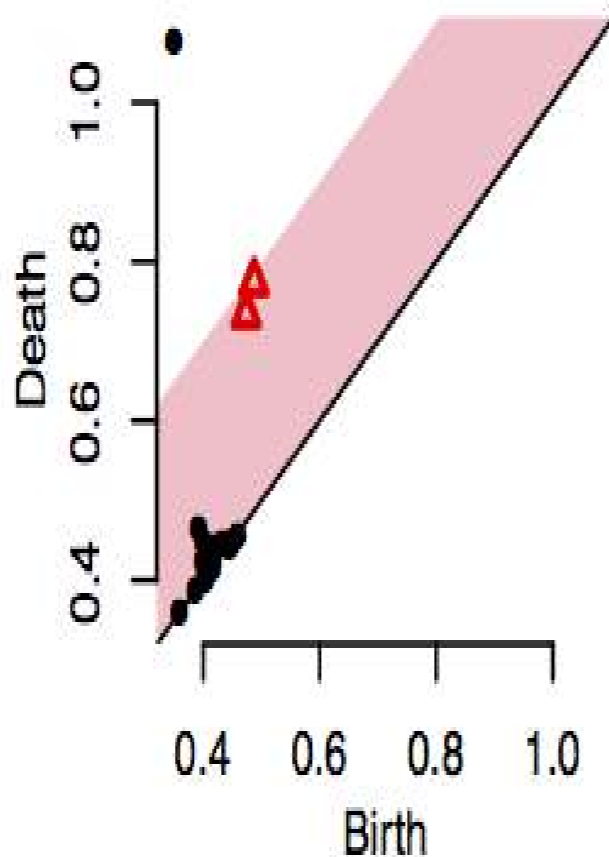
# Bootstrap and significance of topological features

For both methods we can identify significant features by putting a band of size  $2\hat{c}_\alpha$  or  $2\hat{t}_\alpha$  around the diagonal:

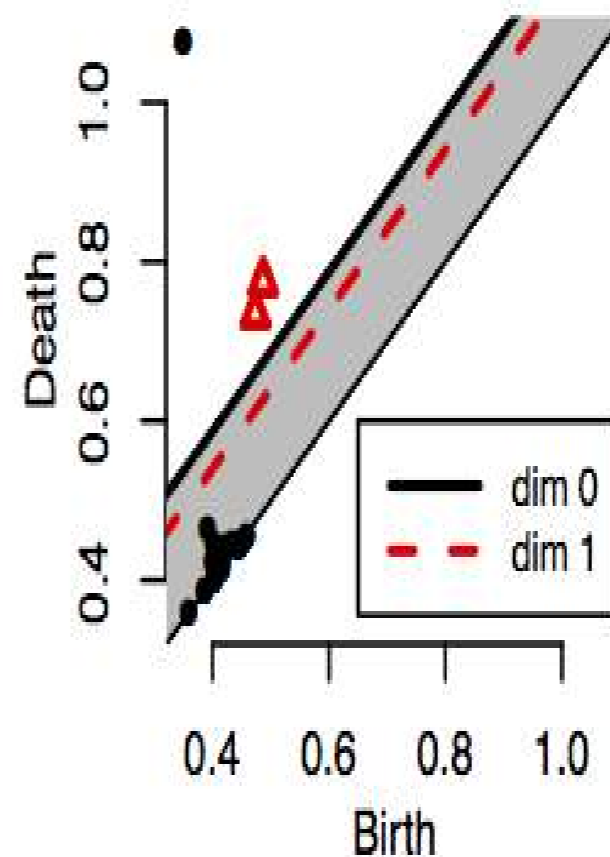
Cassini with Noise



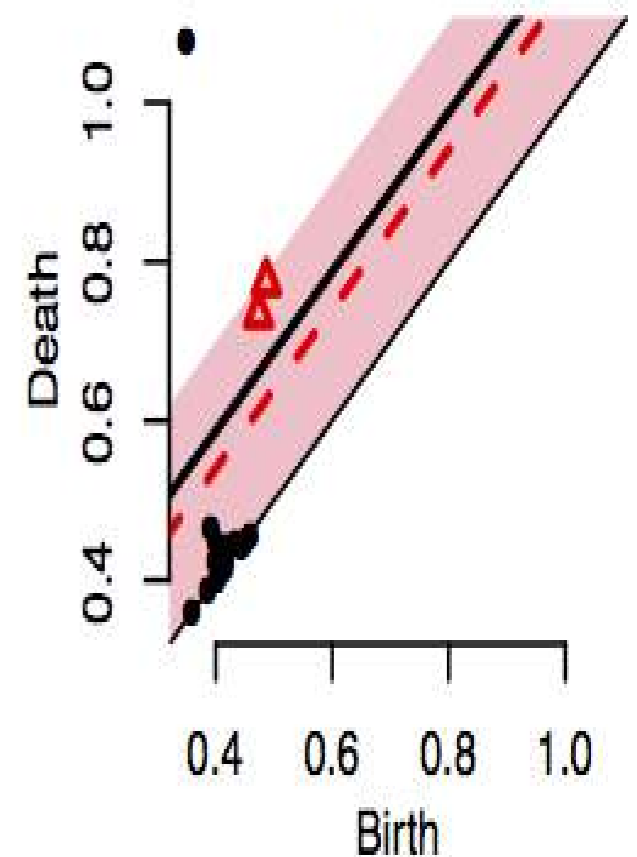
DTM Bootstrap



Bottleneck Bootstrap



Together



In practice, the bottleneck bootstrap can lead to more precise inferences because in many cases the stability result is not sharp enough:

$$d_b(\widehat{Dgm}, Dgm) \leq \|d_{P,m} - d_{P_n,m}\|_\infty.$$



# Concluding remarks

- TDA methods focus on the topological properties (homology / persistent homology) of a shape.
- TDA methods can be used
  - as an “exploratory method”, in particular when the point cloud is sampled on (close to) a real geometric object
  - as a “feature extraction” procedure, next these extracted features can be used for learning purposes.
- TDA is an emerging field, at the interface maths, computer sciences, statistics.
- Many topics about the statistical analysis of TDA
- Applications in many fields of sciences ( medicine, biology, dynamic systems, astronomy, dynamical systems, physics ...)